# Discrete choice (2)

**Applied Econometrics for Spatial Economics** 

## **Hans Koster**

Professor of Urban Economics and Real Estate







- 1. Introduction
- 2. Linear probability model
- 3. Logit
- 4. Probit
- 5. Application
- 6. Summary

- **Topics:** 
  - 1. Discrete choice
    - Random utility framework, estimating binary and multinomial regression models
  - 2. Spatial econometrics
    - Spatial data, autocorrelation, spatial regressions
  - 3. Identification
    - Research design, IV, OLS, RDD, quasi-experiments, standard errors
  - 4. Hedonic pricing
    - Theory and estimation
  - 5. Quantitative spatial economics
    - General equilibrium models in spatial economics



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# Wednesday

09:30-10:30	Lecture 1	Discrete Choice I (The random utility framework)
10:45-11:45	Lecture 2	Discrete Choice II (Estimating discrete choice models)
12:00-13:00	Lecture 3	Spatial Econometrics I (Spatial data)
14:00-15:30	<b>Tutorial 1</b>	Assignment 1

# **Thursday**

09:30-10:30	Lecture 4	Spatial Econometrics II (Spatial autocorrelation)
10:45-11:45	Lecture 5	Spatial Econometrics III (Spatial regressions)
12:00-12:30	Lecture 6	Identification I (Research design)
13:30-14:00	<b>Tutorial 2</b>	Discussion of Assignment 1
14:00-15:00	<b>Tutorial 3</b>	Assignment 2

# Friday

09:30-10:00	Lecture 7	Identification II (RCTs, OLS, IV, quasi-experiments)
10:00-10:30	Lecture 8	Hedonic pricing I (Theory)
10:45-11:45	Lecture 9	Hedonic pricing II (Estimation)
12:00-12:30	<b>Tutorial 4</b>	Discussion of Assignment 2



#### 1. Introduction

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- How to estimate binary discrete choice models?
- Three main options
  - 1. Linear probability model
  - 2. Logit
  - 3. Probit



- 1. Introduction
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# Regress 0/1 variable on characteristics of that choice and use OLS:

$$\Pr(d_j = 1) = \beta' x_j$$

## Dataset example:

Chosen	Price	Time
1	14	12
0	25	5
0	15	15
1	15	13
1	4	45
1	3	40
0	20	10



#### 2. Linear probability model

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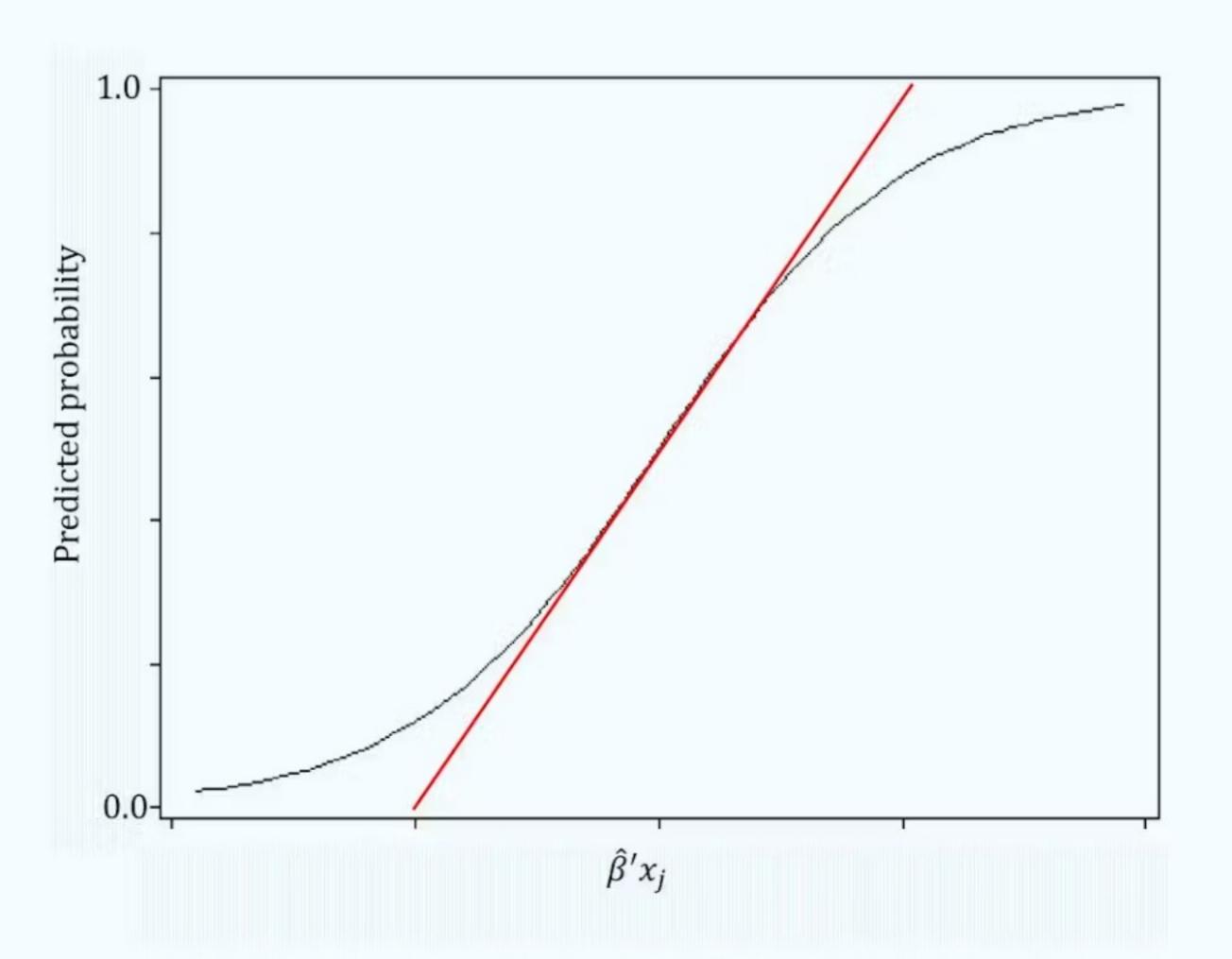
# Advantages:

■ Consistent when  $0 \le \hat{y}_j \le 1 \ \forall j$ 



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■ Consistent when  $0 \le \hat{y}_j \le 1 \ \forall j$ 





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- Consistent when  $0 \le \hat{y}_j \le 1 \ \forall j$
- Easy to interpret
  - Say that  $\beta = -0.25$  and x is measured in  $\mathbb{C}$ , then for each euro increase in x, the probability to choose alternative j decreases by 25 percentage points

$$\frac{\partial \Pr(d_j=1)}{\partial x} = \beta$$



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- Consistent when  $0 \le \hat{y}_j \le 1 \ \forall j$
- Easy to interpret

$$\frac{\partial \Pr(d_j=1)}{\partial x} = \beta$$

- Computationally feasible
  - Important for large panel datasets
- In practice, leads to very similar results as Logit and Probit



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## Disadvantages:

- No direct link with structural parameters of utility function
  - e.g. not able to calculate aggregate utility from choice set

- Biased for small samples and possibly inconsistent marginal effects
  - Linearity?

Not suitable for multinomial choices



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### Let's define

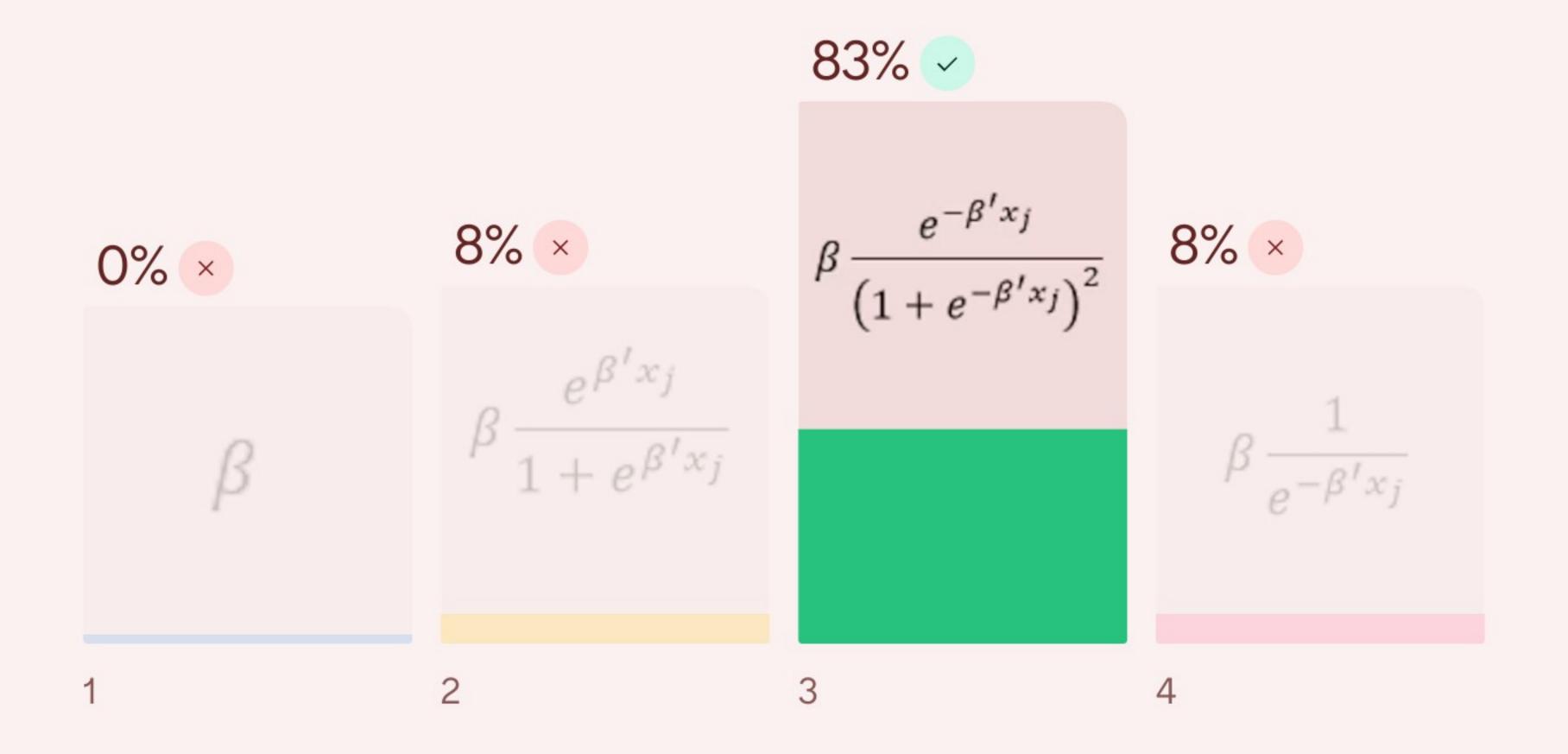
$$\Pr(d_j = 1) = \frac{1}{1 + e^{-\beta' x_j}}$$

# Example: regress 0/1 variable on differences in characteristics of the alternatives

Chosen <sub>B</sub>	Price <sub>B</sub> -Price <sub>A</sub>	Time <sub>B</sub> -Time <sub>A</sub>
1	-14	5
0	5	0
0	15	-20
1	-8	13
1	-10	3
1	3	-5
0	20	10



What is the marginal effect on the probability of one unit increase in x, so  $rac{\partial \Pr(d_j=1)}{\partial x_j}$  ?



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- Marginal effects:
  - Use chain rule of differentiation

• 
$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = -\left(1 + e^{-\beta' x_j}\right)^{-2} \times e^{-\beta' x_j} \times -\beta$$

$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \frac{e^{-\beta' x_j}}{\left(1 + e^{-\beta' x_j}\right)^2}$$



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 The change in the probability for one unit increase in x

$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \frac{e^{-\beta' x_j}}{\left(1 + e^{-\beta' x_j}\right)^2}$$

- Marginal effect depends on x<sub>j</sub>, so is not constant/linear
  - For example, evaluate at mean values of x



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- Software
  - LOGIT or LOGISTIC in STATA
  - REGRESSION BINARY LOGISTIC in SPSS
- In STATA you can select to report marginal effects
  - Use MARGINS after LOGIT command
  - Choose at which x the values are evaluated (e.g. at means)



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## **Advantages of Logit:**

- Predicted probability is always between one and zero
- Clear link to random utility framework
  - Log-sum may be used for welfare calculations
- Closed-form marginal effects
  - Usually leads to very similar results as Probit
- Can include 'fixed effects' (XTLOGIT in STATA)
  - e.g. to control for individual heterogeneity



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## Disadvantages of Logit:

• Why Extreme Value Type I distribution for  $\epsilon$ ?

Maximum likelihood / non linear model



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- We may also assume that  $\epsilon_j$  is normally distributed, so  $\epsilon_j = N(0, \sigma^2)$ 
  - This implies  $Pr(d_i = 1) = \Phi(\beta' x_i)$
  - However, no closed-form for cumulative normal distribution!

Marginal effects:

$$\frac{\partial \Pr(d_j=1)}{\partial x_j} = \beta \phi(\beta x_j)$$

where  $\phi(\cdot)$  is the density function of the normal distribution



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- Normal distribution for  $\epsilon_j$  may seem more reasonable
- Probability is always between one and zero

## Disadvantages:

- No closed-form marginal effects
- Hard to include many fixed effects



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- How to choose between the three models?
  - Probit estimates ≈ Logit estimates
  - Check for robustness of marginal effects
  - Large sample and interested in marginal effects?
    - → Usually linear probability model!
    - → There is an ongoing debate in economics on this issue



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- Koster and Koster (2014)
  - Estimate the value of time and unreliability
  - Uses a stated choice experiment

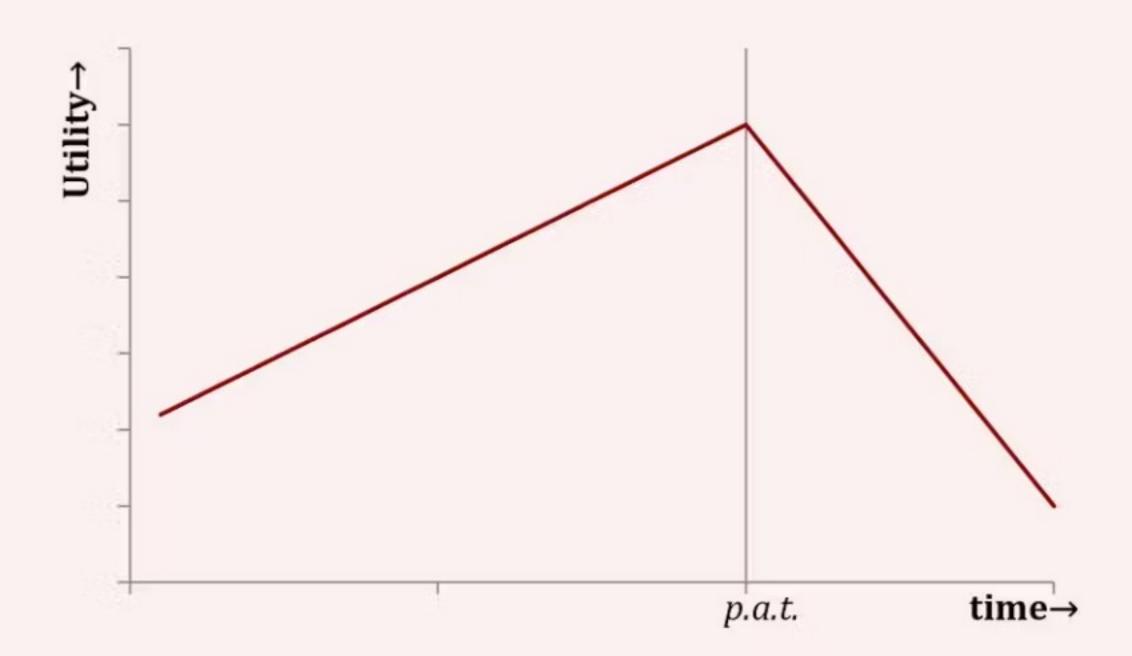
- Stated-choice experiment about preferences of morning commuters
  - "Spitsmijden" (Peak-avoidance project)
  - People get a reward if they avoid the peak
  - But: they may be too early or late at work!
  - Trade-off



#### 5. Application

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- **Value of time**
- Value of schedule delay early
- Value of schedule delay late
  - p.a.t. = preferred arrival time





#### 5. Application

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Example of a choice with two alternatives and uncertainty

Your preferred arrival time if there is no delay is: 8:40.

	Altern	ative 1	Altern	ative 2
Departure time from home	6:	05	6:	50
Probability	80%	20%	90%	10%
Total travel time	30 min	40 min	20 min	35 min
Arrival time at work	6:35	6:45	7:10	7:25
Reward	4 euro	4 euro	0 euro	0 euro



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• 
$$U_{icj} = \beta^R R_{icj} + \beta^T T_{icj} + \beta^{SDE} SDE_{icj} + \beta^{SDL} SDL_{icj} + \epsilon_{icj}$$

individual

c choice

*j* alternative

 $R_{icj}$  expected reward

 $T_{icj}$  expected travel time

 $SDE_{ici}$  expected time before p.a.t.

 $SDL_{ici}$  expected time after p.a.t.

 $\epsilon_{icj}$  random taste variation, Extreme

Value Type I distributed



### 5. Application

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Utility is specified as follows

• 
$$U_{icj} = \beta^R R_{icj} + \beta^T T_{icj} + \beta^{SDE} SDE_{icj} + \beta^{SDL} SDL_{icj} + \epsilon_{icj}$$

• 
$$\Delta u_{inj} = 0 = \beta^R \Delta R_{icj} + \beta^T \Delta T_{icj} + \beta^{SDE} \Delta SDE_{icj} + \beta^{SDL} \Delta SDL_{icj}$$



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## Estimate the value of time and unreliability

Value of time (VOT):

$$-\beta^R \Delta R_{icj} = \beta^T \Delta T_{icj} \rightarrow \Delta T_{icj} = -1 \rightarrow -\Delta R_{icj} = -\frac{\beta^T}{\beta^R}$$

Note that we look at the willingness to <u>pay</u>. Because the experiment focuses on rewards, we have  $-\Delta R_{icj}$ 

Value of schedule delay early (VSDE):

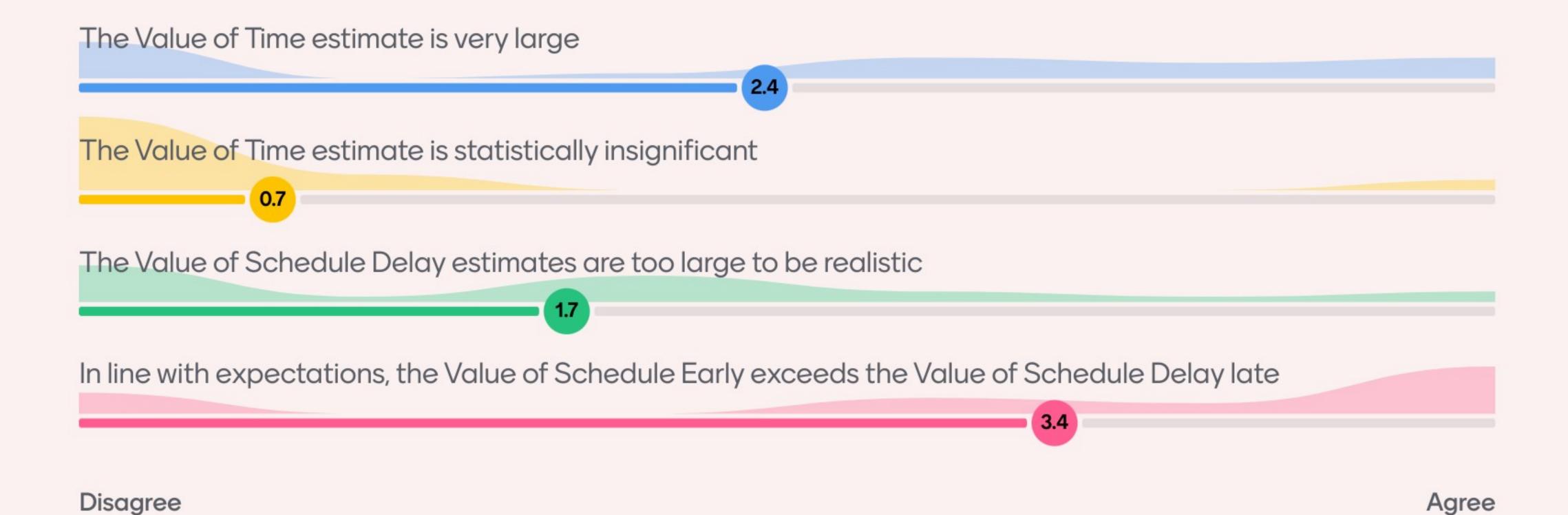
$$-\beta^R \Delta R_{icj} = \beta^{SDE} \Delta SDE_{icj} \rightarrow \Delta SDE_{icj} = -1 \rightarrow -\Delta R_{icj} = -\frac{\beta^{SDE}}{\beta^R}$$

Value of time (VSDL):

$$-\beta^R \Delta R_{icj} = \beta^{SDL} \Delta SDL_{icj} \rightarrow \Delta SDL_{icj} = -1 \rightarrow -\Delta R_{icj} = -\frac{\beta^{SDL}}{\beta^R}$$



# What is your interpretation of the results?



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Results (s.e.'s between parentheses)

VOT	€ 35.05
	(€ 4.158)
VSDE	€ 23.22
	(€ 2.211)
VSDL	€ 17.16
	(€ 1.621)

- Willingness to pay estimates are high
  - People are more sensitive to tolls
  - Relatively high share of high income households
- VSDE>VSDL?
  - Constraints in the morning rather than at work



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## **Today:**

- How to estimate binary choice models?
  - Use LPM, Logit or Probit

 Application to measure value of time, value schedule delay early and schedule delay late



# Discrete choice (2)

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# Discrete choice (3)

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- 1. Introduction
- 2. Multinomial logit

Discrete choice (3)

- 3. Nested logit
- 4. Conditional logit
- 5. RP and SP data
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## Topics:

- 1. Discrete choice
  - Random utility framework, estimating binary and multinomial regression models
- 2. Spatial econometrics
  - Spatial data, autocorrelation, spatial regressions
- 3. Identification
  - Research design, IV, OLS, RDD, quasi-experiments, standard errors
- 4. Hedonic pricing
  - Theory and estimation
- 5. Quantitative spatial economics
  - General equilibrium models in spatial economics



#### Discrete choice (3) 1. Introduction

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## Wednesday

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10:45-11:45	Lecture 2	Discrete Choice II (Estimating discrete choice models)
12:00-13:00	Lecture 3	Spatial Econometrics I (Spatial data)
14:00-15:30	<b>Tutorial 1</b>	Assignment 1

# **Thursday**

09:30-10:30	Lecture 4	Spatial Econometrics II (Spatial autocorrelation)
10:45-11:45	Lecture 5	Spatial Econometrics III (Spatial regressions)
12:00-12:30	Lecture 6	Identification I (Research design)
13:30-14:00	<b>Tutorial 2</b>	Discussion of Assignment 1
14:00-15:00	Tutorial 3	Assignment 2

# Friday

09:30-10:00	Lecture 7	Identification II (RCTs, OLS, IV, quasi-experiments)
10:00-10:30	Lecture 8	Hedonic pricing I (Theory)
10:45-11:45	Lecture 9	Hedonic pricing II (Estimation)
12:00-12:30	<b>Tutorial 4</b>	Discussion of Assignment 2



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## • How to estimate these types of models?

### <u>Overview</u>

	# Alternatives	Coefficients
1. Binary Logit	2	Homogeneous
2. Multinomial Logit with alternative specific parameters	>2, <~10	Differ between alternatives
3. Nested Logit	>2, <~10	Usually homogeneous
4. Conditional Logit	>2	Homogeneous



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#### Recall:

$$Pr(Y = A) = \frac{e^{\beta x_A}}{e^{\beta x_A} + e^{\beta x_B} + e^{\beta x_C}}$$

But now let the coefficients be alternativespecific:

$$Pr(Y = A) = \frac{e^{\beta_A x_A}}{e^{\beta_A x_A} + e^{\beta_B x_B} + e^{\beta_C x_C}}$$

- We cannot identify all the coefficients  $\beta_A$ ,  $\beta_B$ ,  $\beta_C$ , because we compare the results to a reference category
  - » Think of dummies
- Illustration: we can write the probability only in terms of differences with respect to one reference category, e.g.:

$$Pr(Y = A) = \frac{1}{1 + e^{\beta_B x_B - \beta_A x_A} + e^{\beta_C x_C - \beta_A x_A}}$$



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- Independence of irrelevant alternatives
  - Adding an alternative does not affect the relative odds between two other options considered
  - Solution: use Nested Logit
    - → Allows for correlation within nests

- Software
  - NLOGIT in STATA
  - Use Biogeme software
  - Limdep/nlogit



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- Often, the number of alternatives is very large
  - Location choice
  - Route choice
  - Holiday destinations
  - Choice of car
  - Partner choice
  - ...
- With Multinomial Logit this becomes infeasible
  - Unique coefficients for each alternative
  - Not necessary for large choice sets

Conditional Logit:

$$\Pr(d_j = 1) = \frac{e^{\beta' x_j}}{\sum_{k=1}^{J} e^{\beta' x_k}}$$



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- How to deal with large choice sets?
  - Number of observations in your regressions is number of alternatives × respondents

- 1. Model aggregate choices
- 2. Random selection of alternatives
- 3. Estimate count data models (Poisson)



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- 1. Model aggregate choices
- Modelling location choice
  - Focus on aggregate areas (e.g. municipalities)
- Choice of cars
  - Only distinguish between brands

However, lack of detail makes results less credible



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### 2. Random selection of alternatives

- McFadden (1978)
  - Choose a random subset of J alternatives for each choice set, including the chosen option
  - This should not affect the consistency of the estimated parameters
  - Small-sample properties are yet unclear
- How large should J be?

- Applied in many good papers
  - e.g. Bayer et al. (2007, JPE)



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- 3. Estimate count data models
- Estimate Conditional Logit by means of a Poisson model

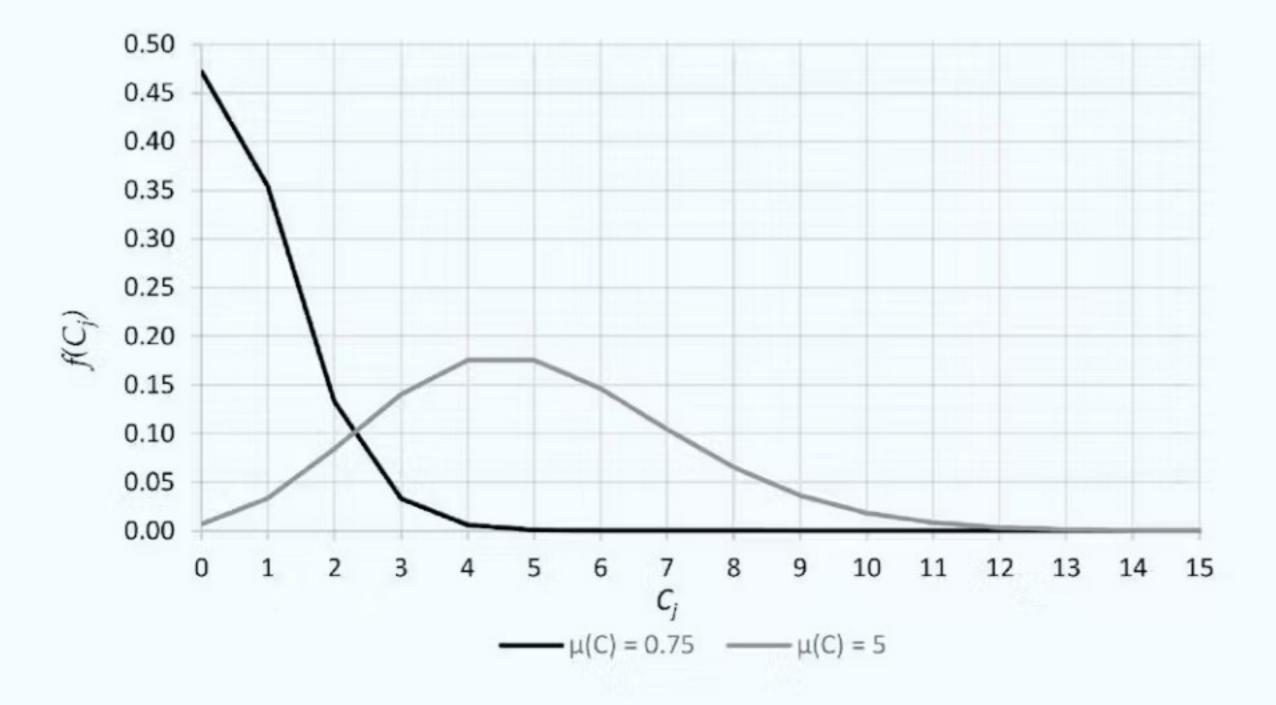
- A <u>Poisson regression</u> is a count data model
  - Dependent variable is integer
  - · ... and should be Poisson distributed



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### 3. Estimate count data models

## Example of a Poisson distribution



Equidispersion:  $\bar{y} = \sigma_y$ 



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- 3. Estimate count data models
- Estimate Conditional Logit by means of a Poisson model

- A <u>Poisson regression</u> is a count data model
  - Dependent variable is integer
  - · ... and should be Poisson distributed
  - $C_j = e^{\beta' x_j} + \epsilon$ where  $C_j$  is the # of decision makers that have chosen a certain alternative

- Convenient interpretation of  $\beta$ 
  - When  $x_j$  increases with one,  $C_j$  increases with  $\beta \times 100$  percent



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- 3. Estimate count data models
- A Poisson model should give identical parameters to the Conditional Logit
  - Maximum likelihood functions are identical up to a constant
  - Guimarães et al. (2003)

- Hence, group observations based on their chosen alternatives
  - ... the number of firms choosing a certain location
  - ... the number of people buying a certain car



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### 3. Estimate count data models

### Implications

- You cannot include characteristics of the decision maker (because you sum over decision makers)!
- Homogeneous parameters across the population

#### Extensions

- Include fixed effects
- Negative binomial regression
- Zero-inflated models
- See Guimarães et al. (2004) for details



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- Jacobs et al. (2013)
  - Analyse location choices of business start-ups
  - Investigate the impact of multinationals on the number of business start-ups
  - In the Randstad Northwing

- Multinationals may generate:
  - Knowledge spillovers
  - Spin-offs
  - Potential customers (output sharing)



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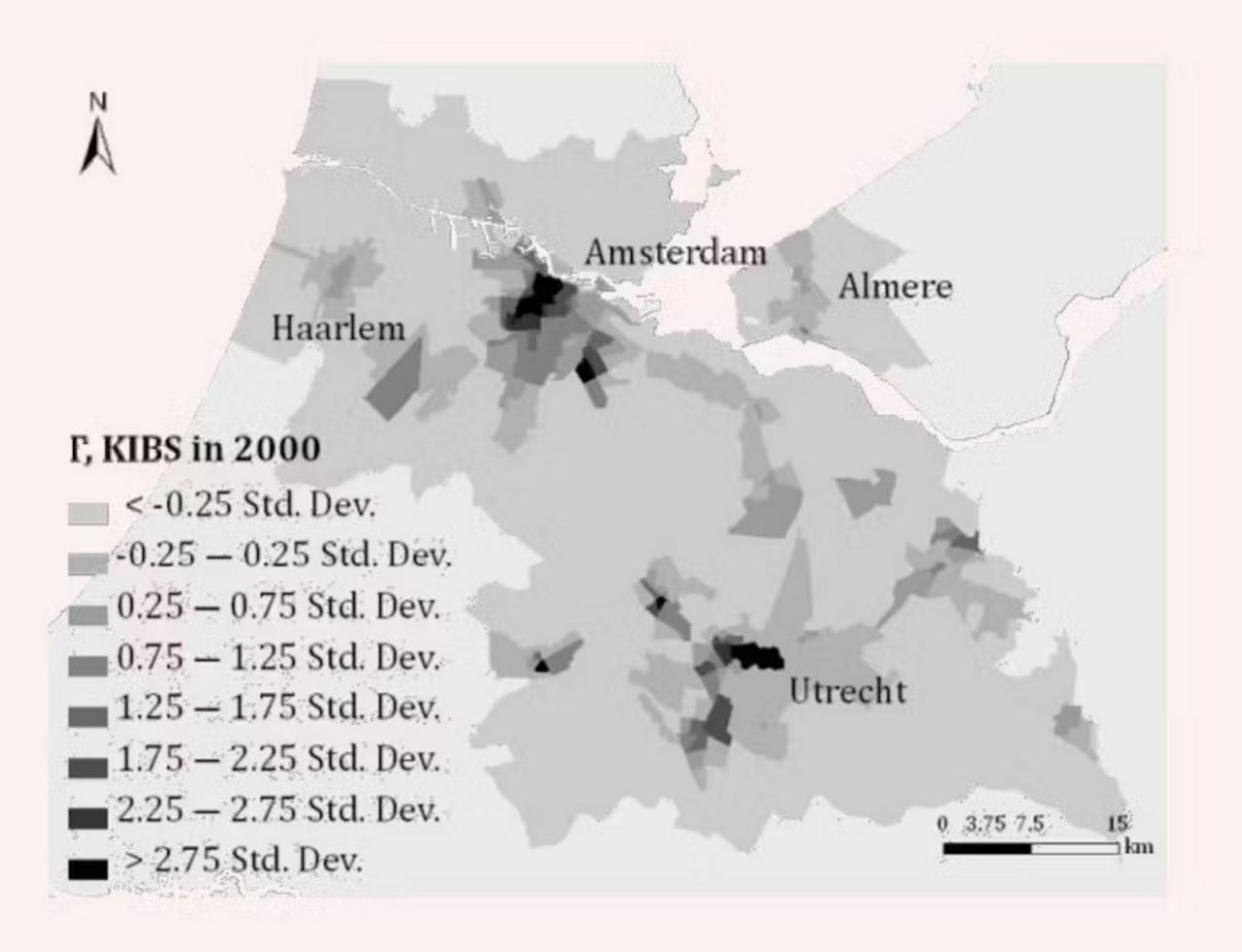
## The Randstad Northwing





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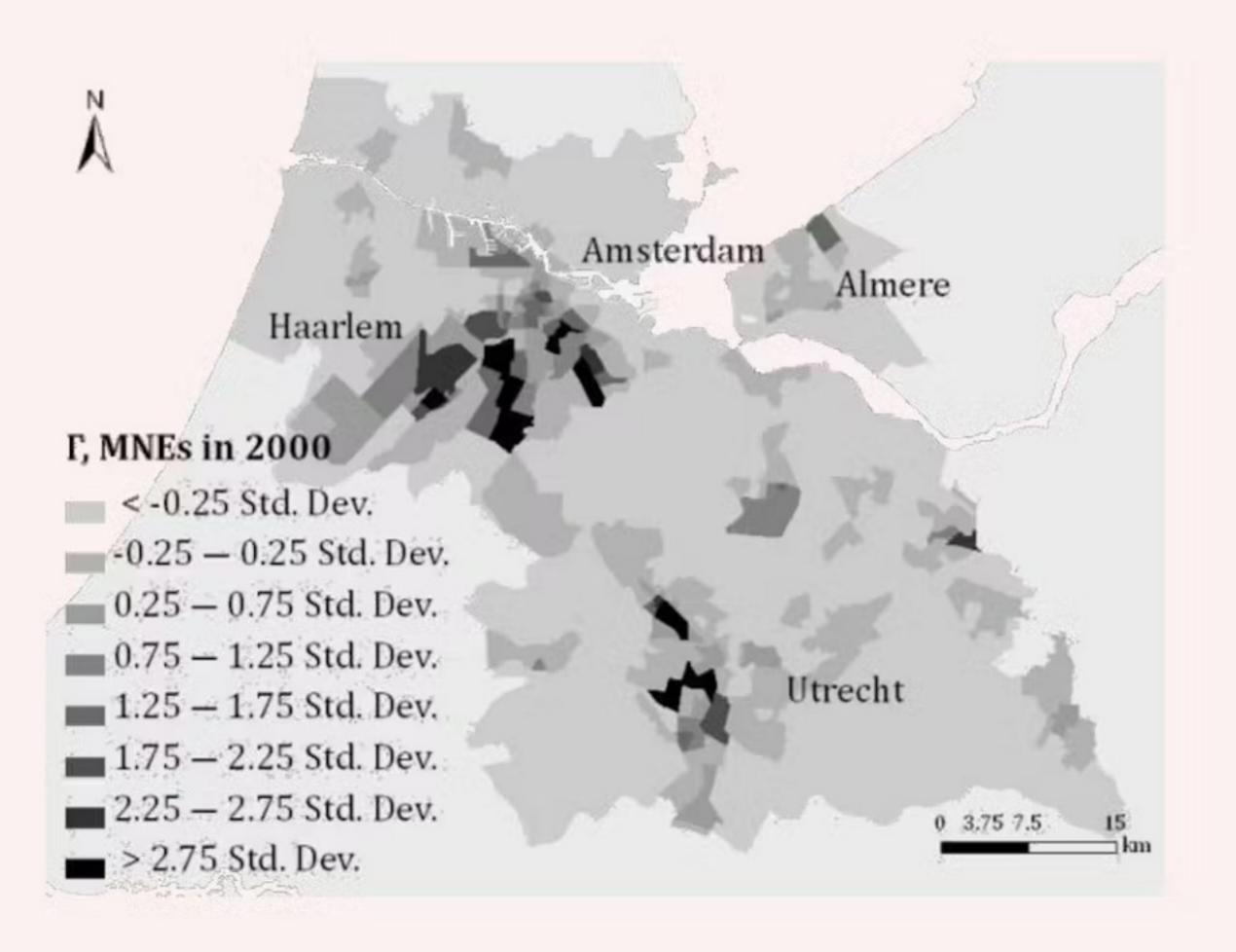
### **Business services in the Randstad Northwing**





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## Multinationals in the Randstad Northwing





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### Location choice model:

$$\pi_{ij} = \alpha + \beta e_j^{MNE} + \gamma e_j^{BSF} + \delta e_j^{OF} + \zeta X_j + \eta_{j \in M} + \epsilon_{ij}$$
 $i$  firm
 $j$  PC6 location (alternatives)
 $e_j^{MNE}$  multinational employment
 $e_j^{BSF}$  business services employment
 $e_j^{OF}$  other employment
 $X_j$  control variables
 $\eta_{j \in M}$  municipality fixed effects



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• Probability  $\Pi$  that i chooses k:

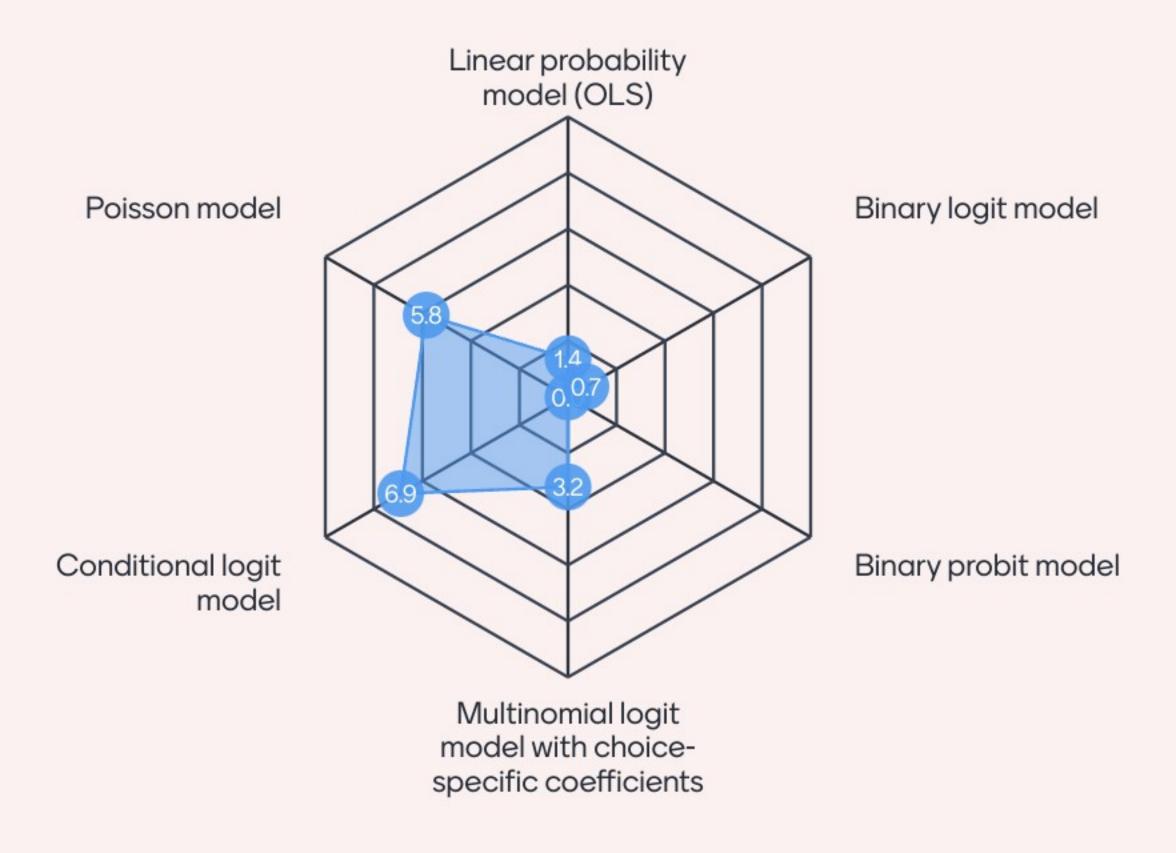
$$\Pr(d_j = 1) = \frac{e^{\alpha + \beta e_j^{MNE} + \gamma e_j^{BSF} + \delta e_j^{OF} + \zeta X_j + \eta_{j \in M}}}{\sum_{\kappa=1}^{J} e^{\alpha + \beta e_k^{MNE} + \gamma e_k^{BSF} + \delta e_k^{OF} + \zeta X_k + \eta_{k \in M}}}$$

There are 13,655 locations

→ How would you estimate this model?



## What regression method would you use to estimate this model?



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- Problem: many locations
  - Use count data to estimate this model
    - → There are no individual firm characteristics
  - Dependent variable: # start-ups per location



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### Results

Table – A POISSON MODEL
(Dependent variable: The number of business services start-ups per location)

	(1)	(2)	(3)
Multi-national employment density (log)	0.0709***	0.0422***	0.0772***
	(0.0151)	(0.0092)	(0.0121)
Business services employment density (log)	0.4304***	0.4374***	0.3821***
	(0.0240)	(0.0162)	(0.0214)
Other employment density (log)	-0.2242***	-0.2203***	-0.1352***
	(0.0162)	0.0071	(0.0178)
Control variables (9)	No	Yes	Yes
Municipality fixed effects (61)	No	No	Yes
Number of locations	13,655	13,655	13,655
Log-likelihood	-13,146.903	-13,051.163	-12,709.249

Notes: We include locations with at least 10 employees in 2000. The coefficients can be interpreted as elasticities and differ from Jacobs et al. (2013) because of a slightly different set of controls and because we estimate Poisson models instead of Negative-binomial regressions. Robust standard errors are in parentheses. \*\*\* p < 0.01, \*\* p < 0.05, \* p < 0.10.

## → Please interpret the results in column (3)



- 1. Introduction
- 2. Multinomial logit
- 3. Nested logit
- 4. Conditional logit
- 5. RP and SP data
- 6. Summary

- Multinationals attract new business services
  - The effect of other business services on startups, is however, much larger

- Coefficients are convenient to interpret
  - e.g. a 1% increase in multinational empl. leads to an increase of start-ups of 0.077% (in Column (1))



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### Discrete choice:

- Random utility framework
- Generalisations of logit models
  - LPM (J = 2)
  - Binary logit/probit (J = 2)
  - Multinomial logit (2 < J < 10)
  - Nested logit (2 < J < 10)
  - Conditional logit (l > 2)
- Conditional Logit models can be estimated by count data models
  - Cannot include characteristics of the decision maker



# Discrete choice (3)

**Applied Econometrics for Spatial Economics** 

## **Hans Koster**

Professor of Urban Economics and Real Estate







- 1. Introduction
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## Types of data

- Revealed preference (RP) data
  - Observed or reported actual behaviour

- Stated preference (SP) data
  - Respondents are confronted with hypothetical choice sets

Combinations of RP and SP



### 5. Revealed preference and stated preference data

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## Advantages of RP data

- Based on actual behaviour!!
- Use existing (large) data sources
  - Cheaper
  - No expensive experiments
- Panels of the same individuals over a long time



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- Lack of variability
- Collinearity (e.g. price and travel times)
- Lack of knowledge on the choice set
- Not possible with new choice alternatives
- Actual behaviour may not be first choice
  - University numerus fixus
- Perception errors and imperfect information
  - Airline tickets



### 5. Revealed preference and stated preference data

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- Example of stated preference question
  - Different from contingent valuation!

Suppose you have to ship a product from A to B					
Option 1		Option 2			
Price:	€ 1,000	Price:	€ 750		
Handling time:	3 days	Handling time:	1 week		
% does not arrive	e: 1.0%	% does not arrive	e: 1.3%		
What alternative will you choose?					



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## Advantages of SP data

- New alternatives
- New attributes
- Large variability is possible
- Problems of collinearity can be solved
  - 'Orthogonal design'
- Choice set is clearly defined



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- Information bias
- Starting point bias
- Hypothetical bias
- Strategic bias
- Errors



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- Information bias
  - The respondent has incorrect information on the context
  - Make your experiment as realistic as possible

- Starting point bias
  - Respondents are influenced by the set of available responses to the experiment
  - Test your design and choose realistic attribute values



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- Hypothetical bias
  - Individuals tend to respond differently to hypothetical scenarios than they do to the same scenarios in the real world.
  - Cognitive incongruity with actual behaviour
  - Again: make your experiment as realistic as possible
  - But otherwise hard to mitigate...

### Strategic bias

- Respondent wants a specific outcome
- (S)he fills in answers that are in line with desired outcomes



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- Unintentional biases
  - Information, starting point, hypothetical bias
- Intentional biases
  - Strategic bias

- Errors
  - Boredom
  - Respondents do not carefully read instructions
  - Respondents do not understand the questions



If there is good data available, I would prefer RP (personal opinion)

## Would you use SP or RP in the following cases:

Investigate the preference of households for hydrogen-powered cars

Test the impact of a (past) industrial policy on firm location choices

Measure costs and benefits of the construction of a high-speed rail line to Shikoku

Measure the economic impact of existing power plants on nearby residential neighbourhoods

0.0

Stated preference

Revealed preference

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## **Today:**

- Generalisations of logit models
  - Multinomial logit
  - Nested logit
  - Conditional logit
- Conditional Logit models can be estimated by count data models
  - Cannot include characteristics of the decision maker
- Data
  - Stated preference or revealed preference data



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