

Spatial econometrics (2)

Applied Econometrics for Spatial Economics

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- 1. [Introduction](#)
- 2. [Spatial autocorrelation](#)
- 3. [Spatial regressions](#)
- 4. [Summary](#)

- **Topics:**

- 1. **Discrete choice**

- Random utility framework, estimating binary and multinomial regression models

- 2. **Spatial econometrics**

- Spatial data, autocorrelation, spatial regressions

- 3. **Identification**

- Research design, IV, OLS, RDD, quasi-experiments, standard errors

- 4. **Hedonic pricing**

- Theory and estimation

- 5. **Quantitative spatial economics**

- General equilibrium models in spatial economics

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Wednesday

09:30-10:30	Lecture 1	Discrete Choice I (The random utility framework)
10:45-11:45	Lecture 2	Discrete Choice II (Estimating discrete choice models)
12:00-13:00	Lecture 3	Spatial Econometrics I (Spatial data)
14:00-15:30	Tutorial 1	Assignment 1

Thursday

09:30-10:30	Lecture 4	Spatial Econometrics II (Spatial autocorrelation)
10:45-11:45	Lecture 5	Spatial Econometrics III (Spatial regressions)
12:00-12:30	Lecture 6	Identification I (Research design)
13:30-14:00	Tutorial 2	Discussion of Assignment 1
14:00-15:00	Tutorial 3	Assignment 2

Friday

09:30-10:00	Lecture 7	Identification II (RCTs, OLS, IV, quasi-experiments)
10:00-10:30	Lecture 8	Hedonic pricing I (Theory)
10:45-11:45	Lecture 9	Hedonic pricing II (Estimation)
12:00-12:30	Tutorial 4	Discussion of Assignment 2

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- **Spatial autocorrelation between values**
 - Implies $\text{cov}(x_i, x_j) = E[x_i x_j] - E[x_i] \cdot E[x_j] \neq 0$
 - Again, j refers to other locations
- **Spatial autocorrelation, dependence, clustering**
 - Fuzzy definitions in literature

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- How to measure spatial autocorrelation
 - Moran's I
 - Focus on one variable x (e.g. crime)
- H_0 : independence, spatial randomness
- H_A : dependence
 - On the basis of adjacency, distance, hierarchy

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- Moran's I is given by:

$$I = \frac{R}{S_0} \times \frac{\tilde{x}' W \tilde{x}}{\tilde{x}' \tilde{x}} \quad (4)$$

where R is the number of spatial units
 S_0 is the sum of all elements of the spatial weight matrix
 W is the spatial weight matrix
 $\tilde{x} = x - \bar{x}$ is a vector with the variable of interest

- Use row-standardised spatial weight matrix W !
 - So that $I_S = \frac{\tilde{x}' W \tilde{x}}{\tilde{x}' \tilde{x}}$

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- **Moran's I**
- **Sidenote:**
 - Please realise that $W\tilde{x}$ is a vector
 - $I_S = \frac{\tilde{x}'W\tilde{x}}{\tilde{x}'\tilde{x}}$
 - $\mathbf{W} \times \tilde{\mathbf{x}} = \mathbf{W}\tilde{\mathbf{x}}$
$$\begin{bmatrix} 0 & w_{12} & w_{13} \\ w_{21} & 0 & w_{23} \\ w_{31} & w_{32} & 0 \end{bmatrix} \times \begin{bmatrix} \tilde{x}_1 \\ \tilde{x}_2 \\ \tilde{x}_3 \end{bmatrix} = \begin{bmatrix} \dots \\ \dots \\ \dots \end{bmatrix}$$
 - Notation: $\frac{x'y}{x'x} = \mathbf{x}^T \mathbf{y} (\mathbf{x}^T \mathbf{x})^{-1}$

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- **Moran's I**
- **Recall that $I_S = \frac{\tilde{x}'W\tilde{x}}{\tilde{x}'\tilde{x}}$ (standardised I)**
 - **Note similarity with OLS:** $\hat{\beta} = \frac{x'y}{x'x}$
 - **Hence:** $W\tilde{x} = \alpha + I\tilde{x} + \epsilon$, where $\alpha = 0$
- **Moran's I is correlation coefficient (more or less)**
 - $\approx [-1,1]$
 - **But: expectation** $E[I] = -\frac{1}{N-1}$
- **Visualisation**
 - **Moran scatterplot**

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- Moran's I
- How to investigate the statistical significance of (4)?
 - $\frac{I - E[I]}{\sqrt{\text{var}[I]}}$ (5)
 - However, $\sqrt{\text{var}[I]}$ is difficult to derive
 - $E[I] = -1/(n - 1)$
 - Assume normal distribution of I to approximate $\sqrt{\text{var}[I]}$ under H_0
 - Or: bootstrapping/simulation
 - See Cliff and Ord (1973) for more details

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- Moran's *I*
- Also possible: correlation to other variables:

$$I_S = \frac{\tilde{x}' W \tilde{z}}{\tilde{x}' \tilde{x}}$$

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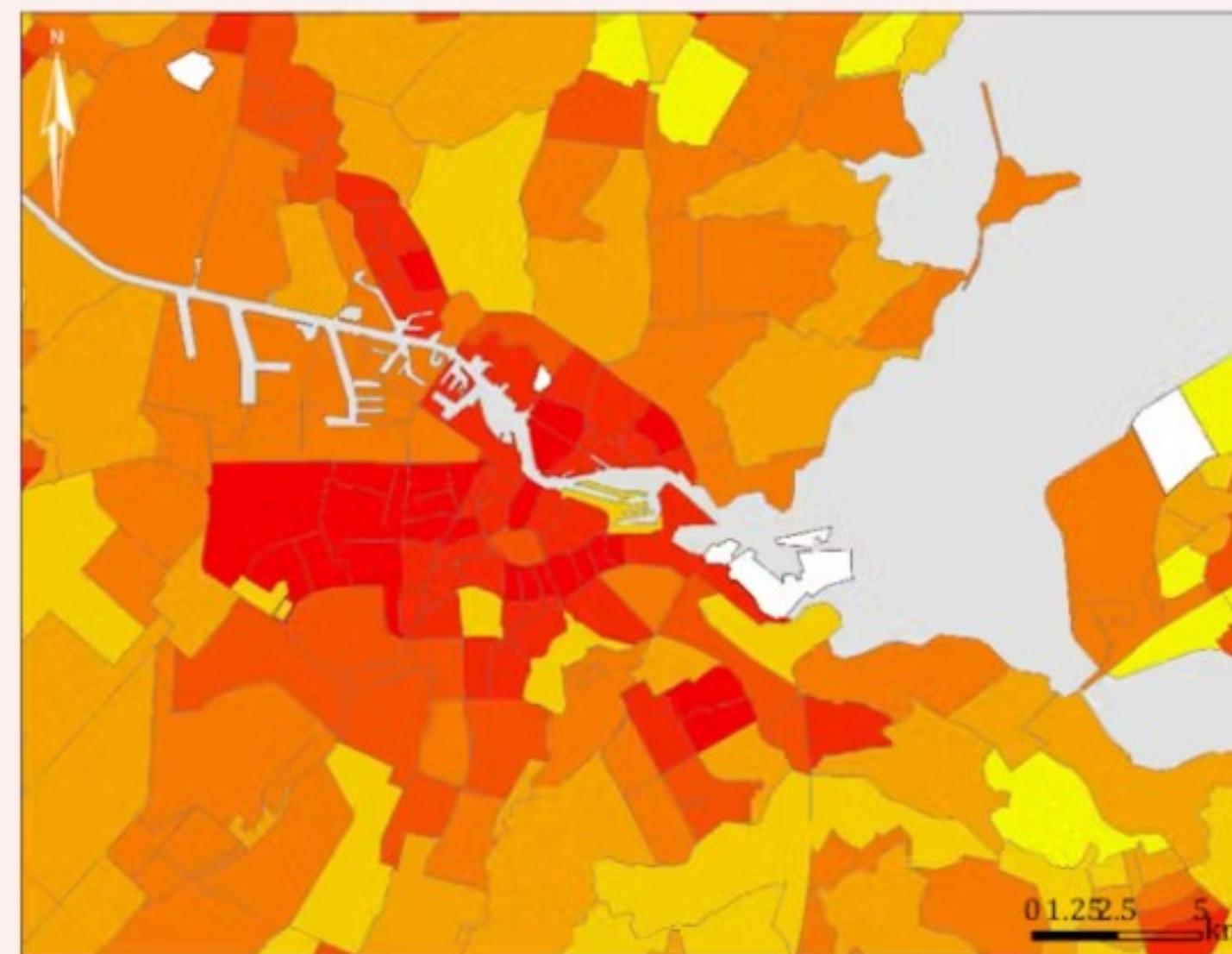
- **How to calculate Moran's I using software**
 - SPAUTOC **in STATA**
 - SPLAGVAR **in STATA**
 - SPATIAL STATISTICS TOOLBOX **in ArcGIS**
- **Alternative: Getis and Ord's G**
 - **Most of the time only Moran's I is reported**

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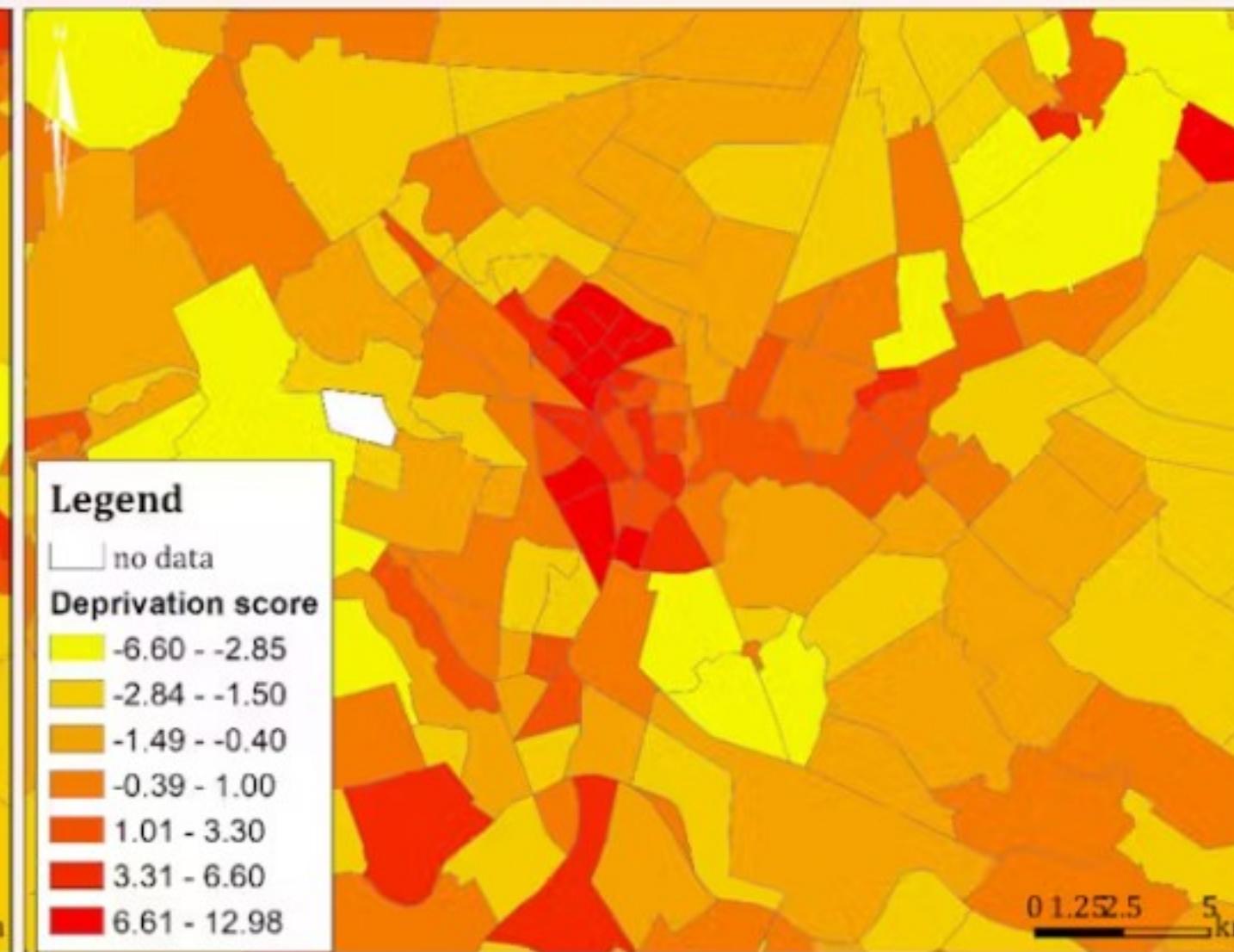
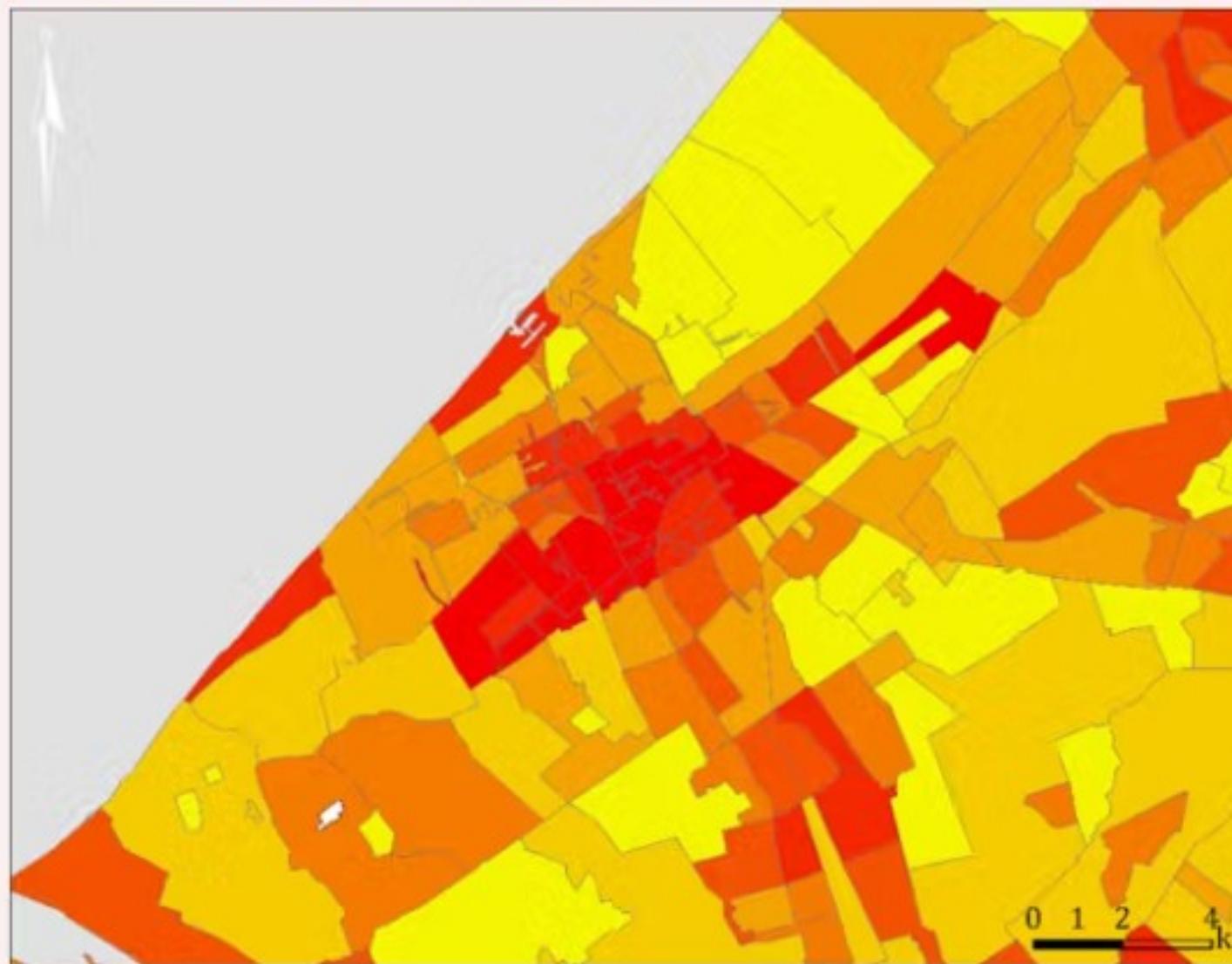
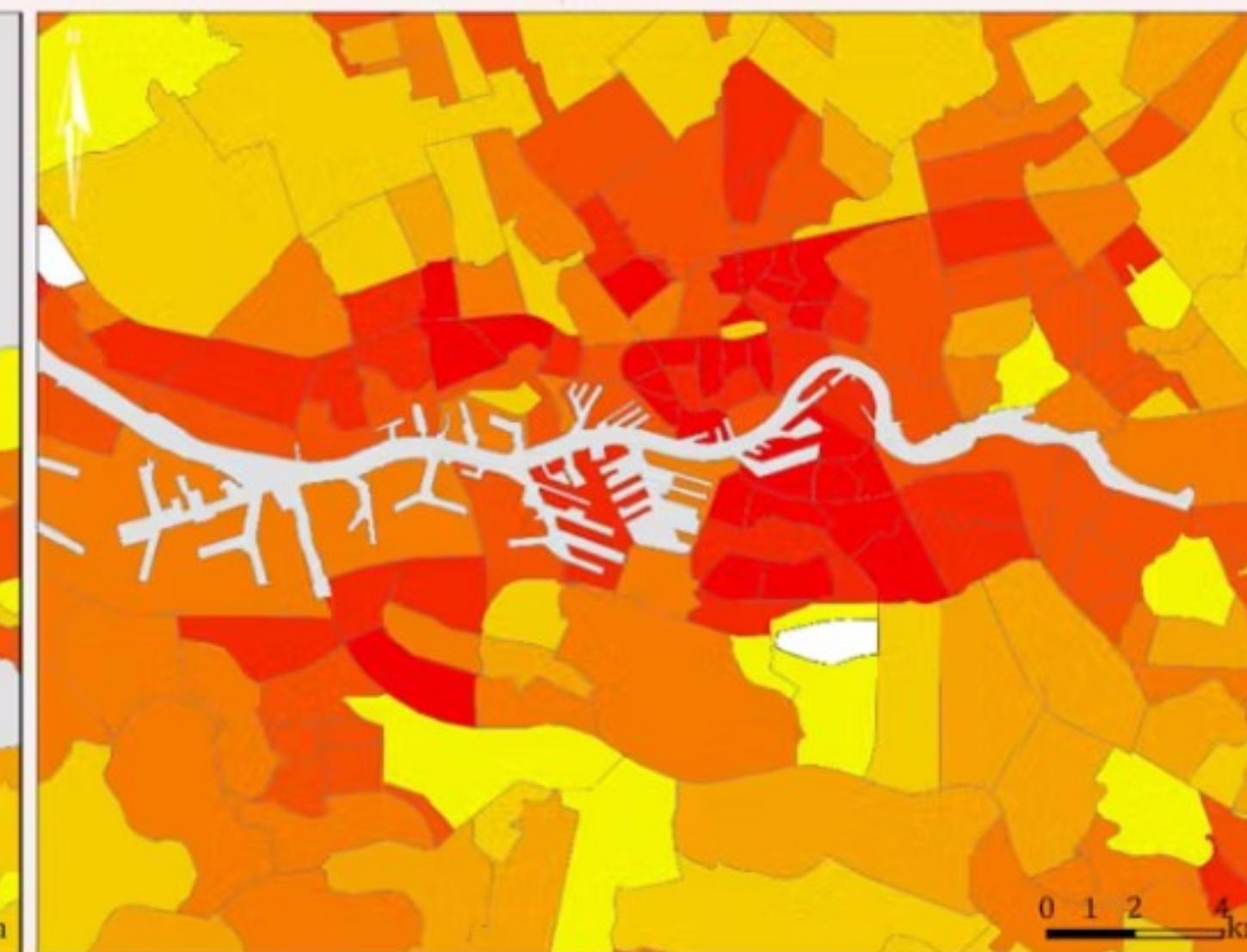
- Let's try to answer the question:
"Is social deprivation spatially clustered?"
- How to determine the most deprived neighbourhoods?
- Dutch government calculated deprivation z-score for each neighbourhood
 - Based on housing quality, safety, perception and satisfaction
 - *Important:* the 83 most deprived neighbourhoods were selected for an investment of >€1 billion

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Amsterdam



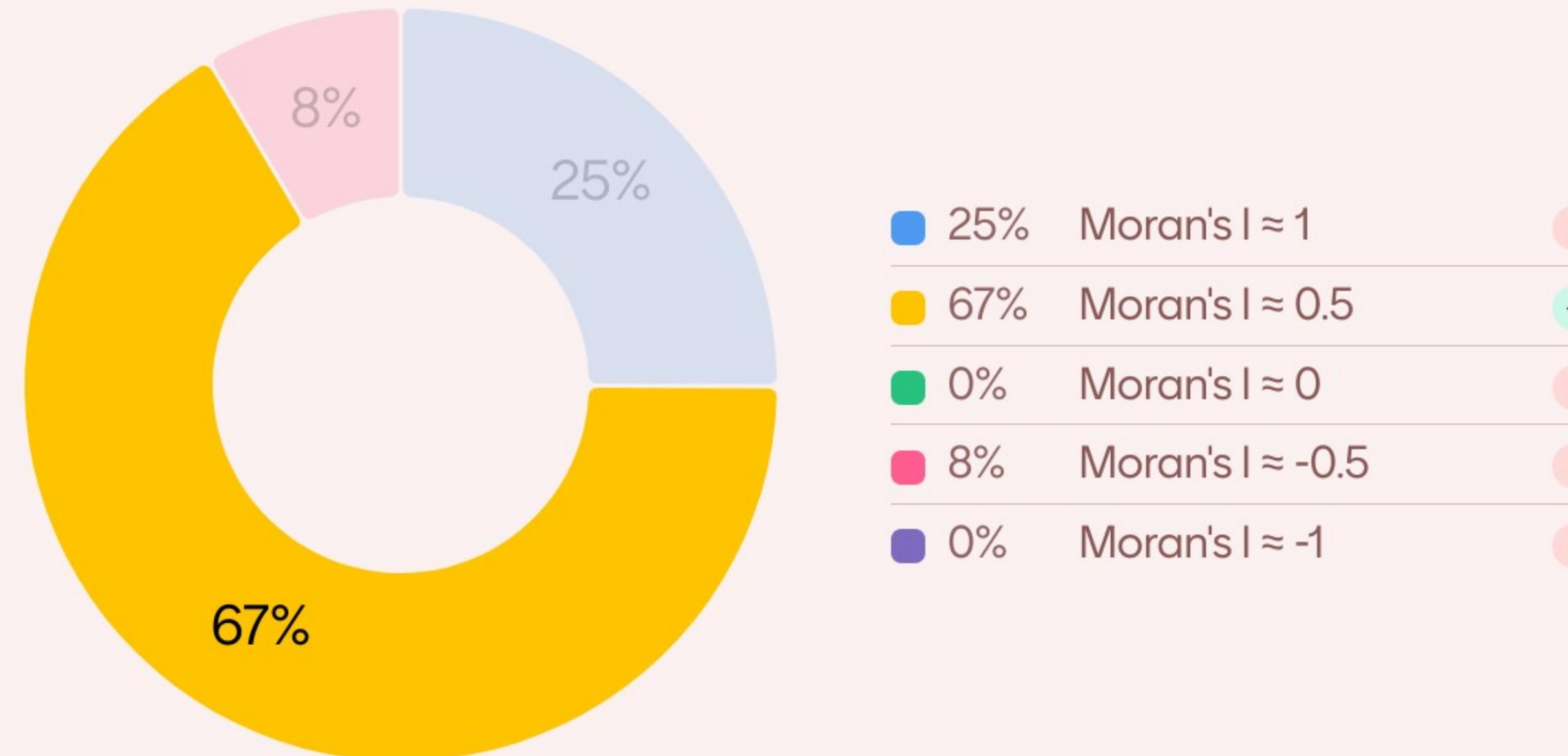
Rotterdam



The Hague

Utrecht

What is your hypothesis when looking at the figure



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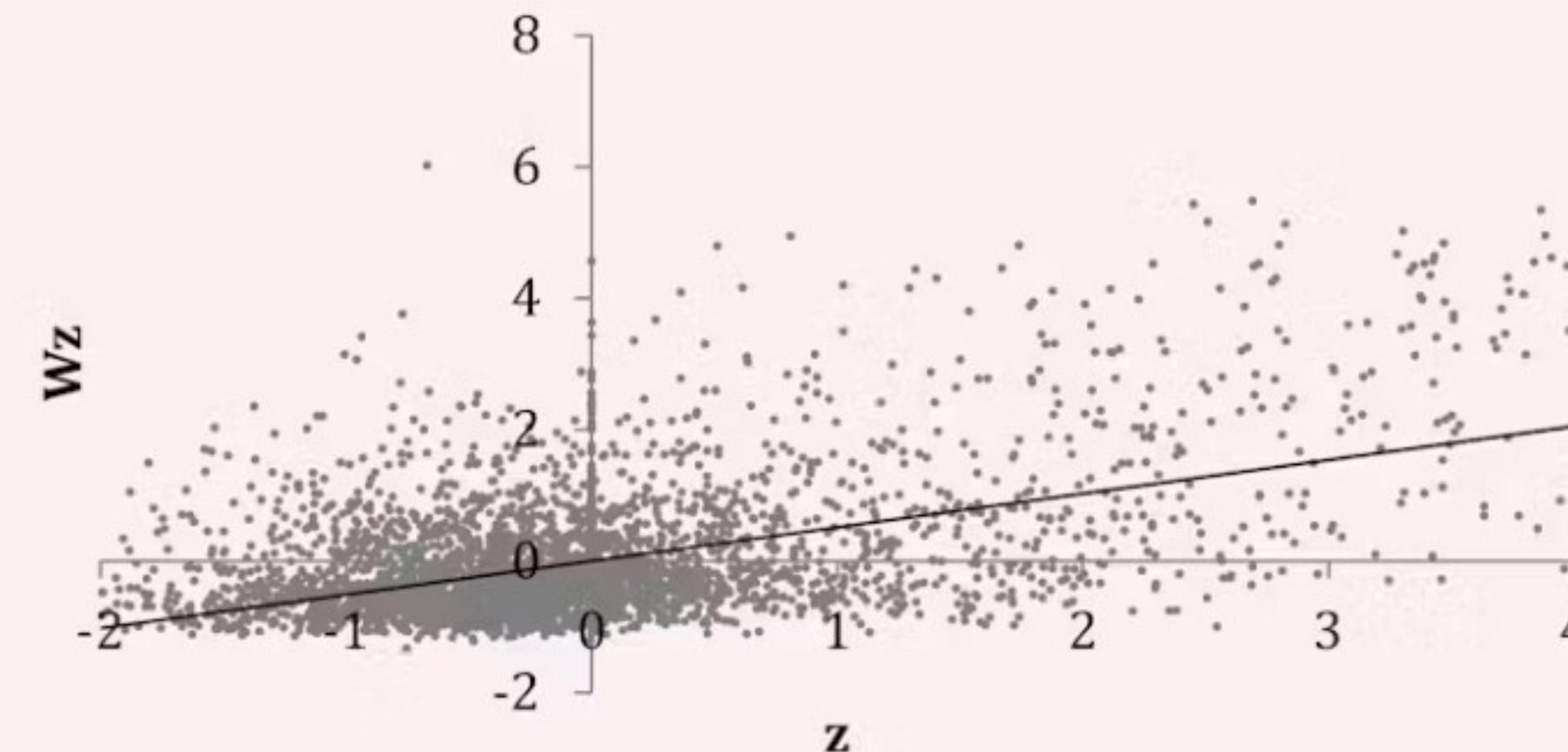
- **Determine spatial autocorrelation**
 1. Determine distance between all neighbourhoods using centroids
 2. Use inverse distance function $w_{ij} = 1/(d_{ij}^\gamma)$ to determine spatial weights in weight matrix
 3. Calculate Moran's I: $W\tilde{\mathbf{z}} = \alpha + I\tilde{\mathbf{z}} + \epsilon$ where $\tilde{\mathbf{z}} = \mathbf{z} - \bar{\mathbf{z}}$ and W is a row-standardised weight matrix
 - *Recall that $W\tilde{\mathbf{z}}$ is a vector*
 4. Bootstrap this procedure to estimate standard error (or use software)

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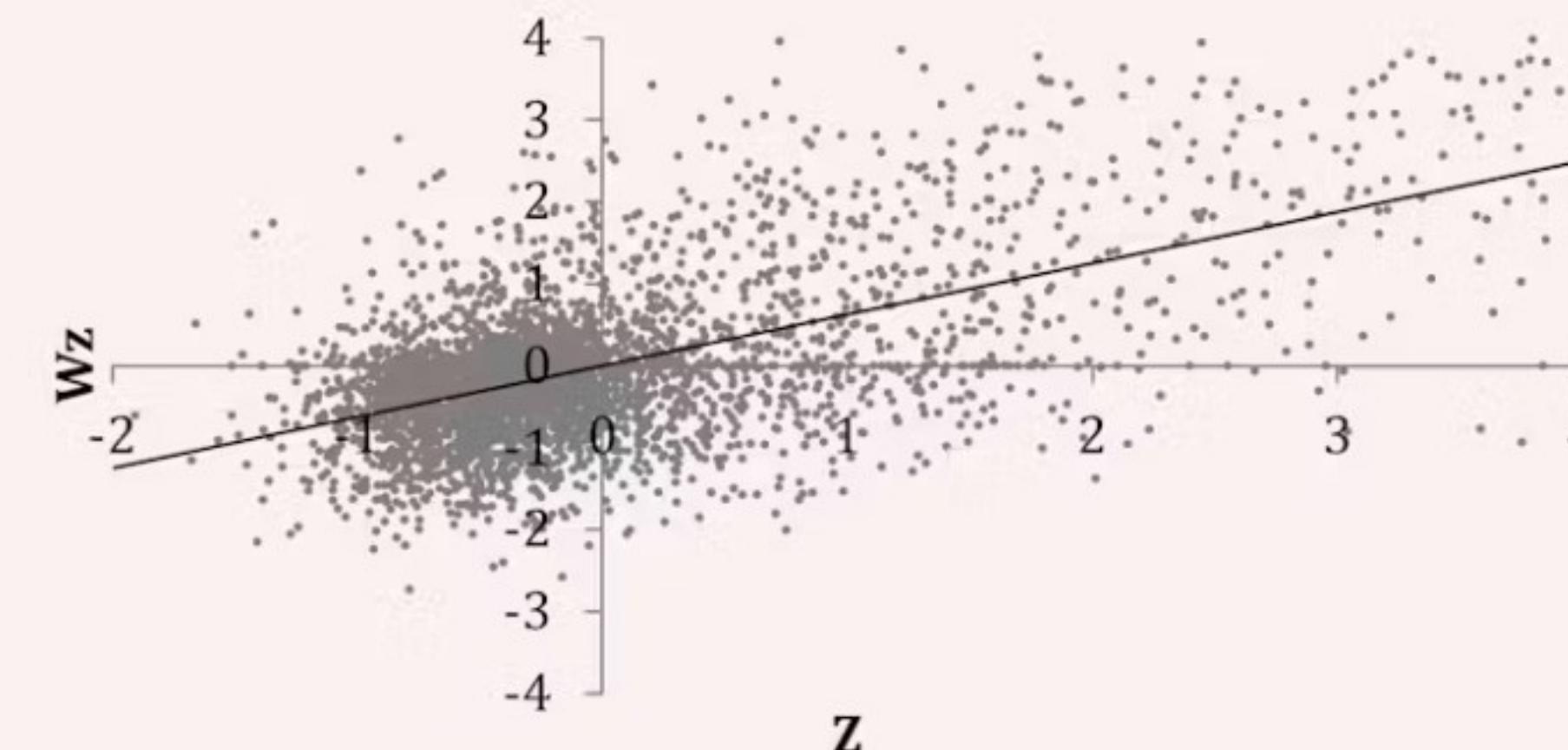
- Calculate Moran's I

- Using inverse distance function $w_{ij} = \frac{1}{d_{ij}^\gamma}$

$$I = 0.513^{***} (\gamma=1)$$



$$I = 0.625^{***} (\gamma=2)$$



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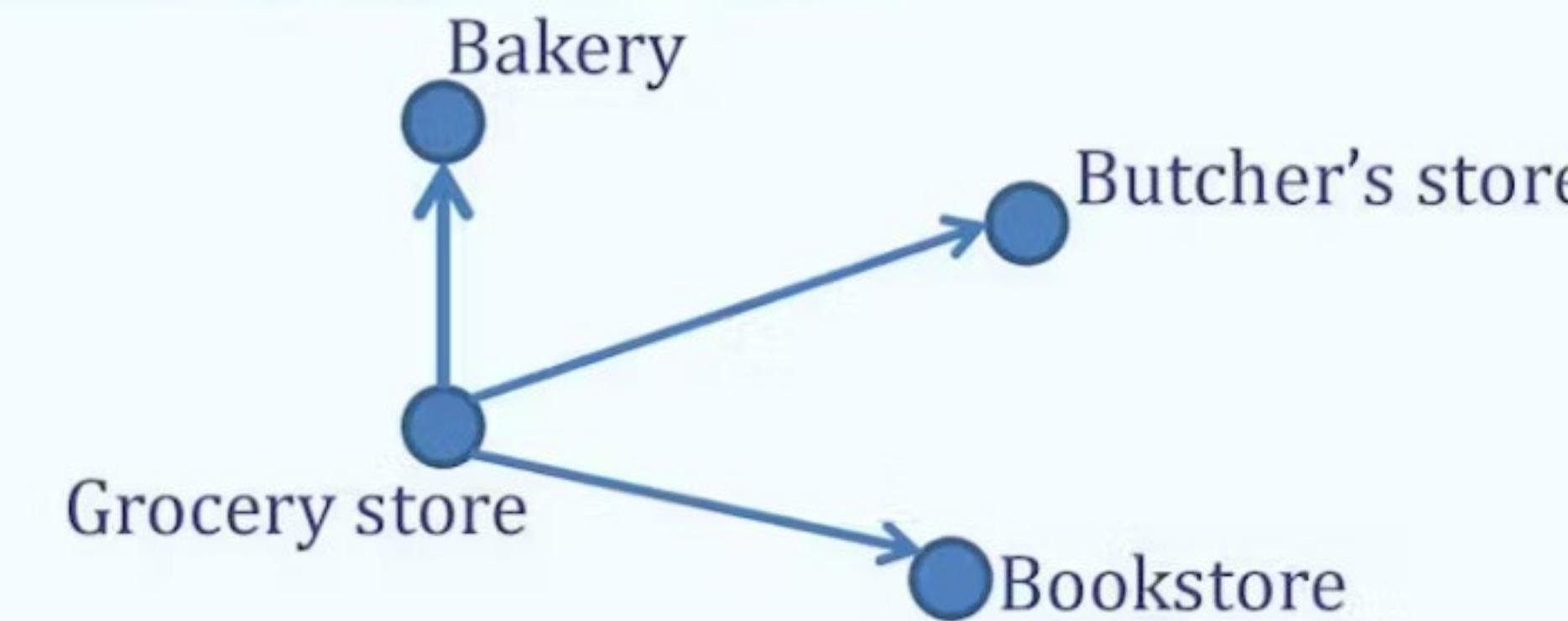
- **Spatial correlation in deprivation**
 - Local phenomenon?
 - You do not know *why* scores are autocorrelated...
 - No causal relationships!

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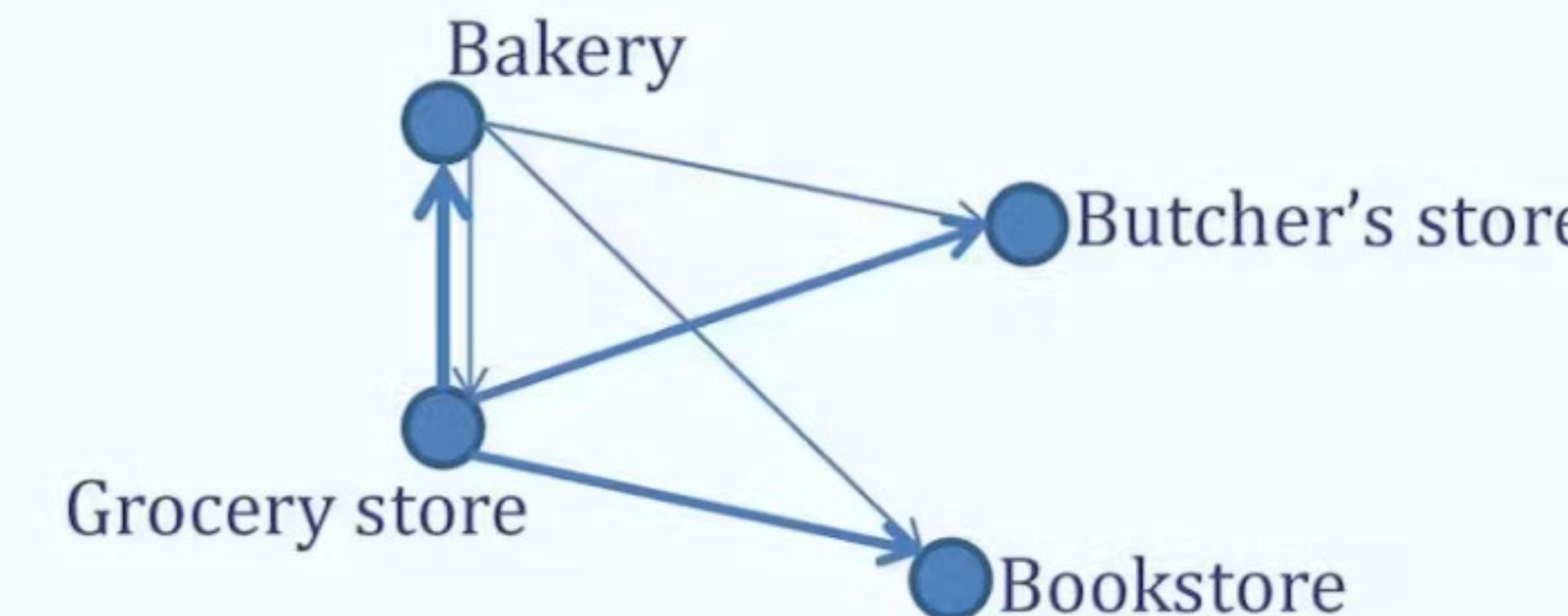
- It is important to make a distinction between *global* and *local* spatial autocorrelation
 - See Anselin (2003) for a discussion
- Global spatial autocorrelation
 - Local shock affects the whole system
- Local spatial autocorrelation
 - Local shock only affects the ‘neighbours’

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- Example: Consider an income increase for grocery store owner
- Local autocorrelation:

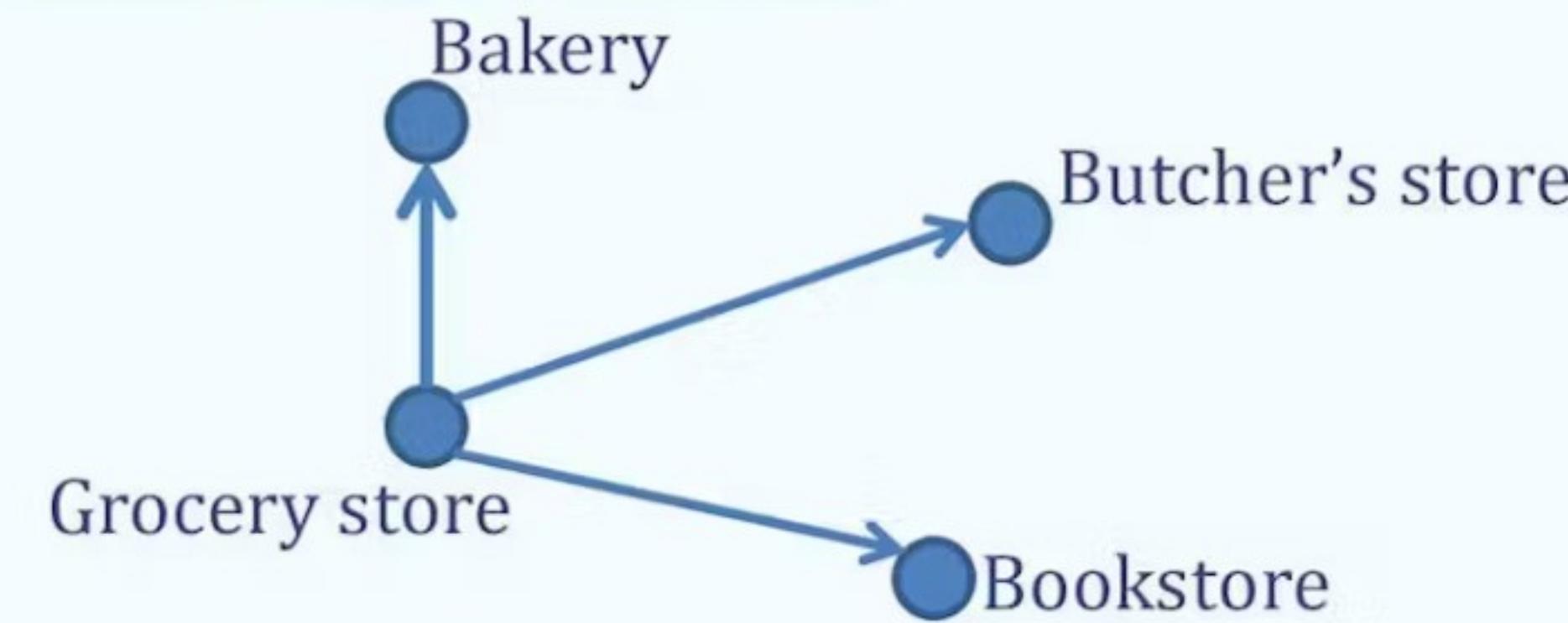


- Global autocorrelation:

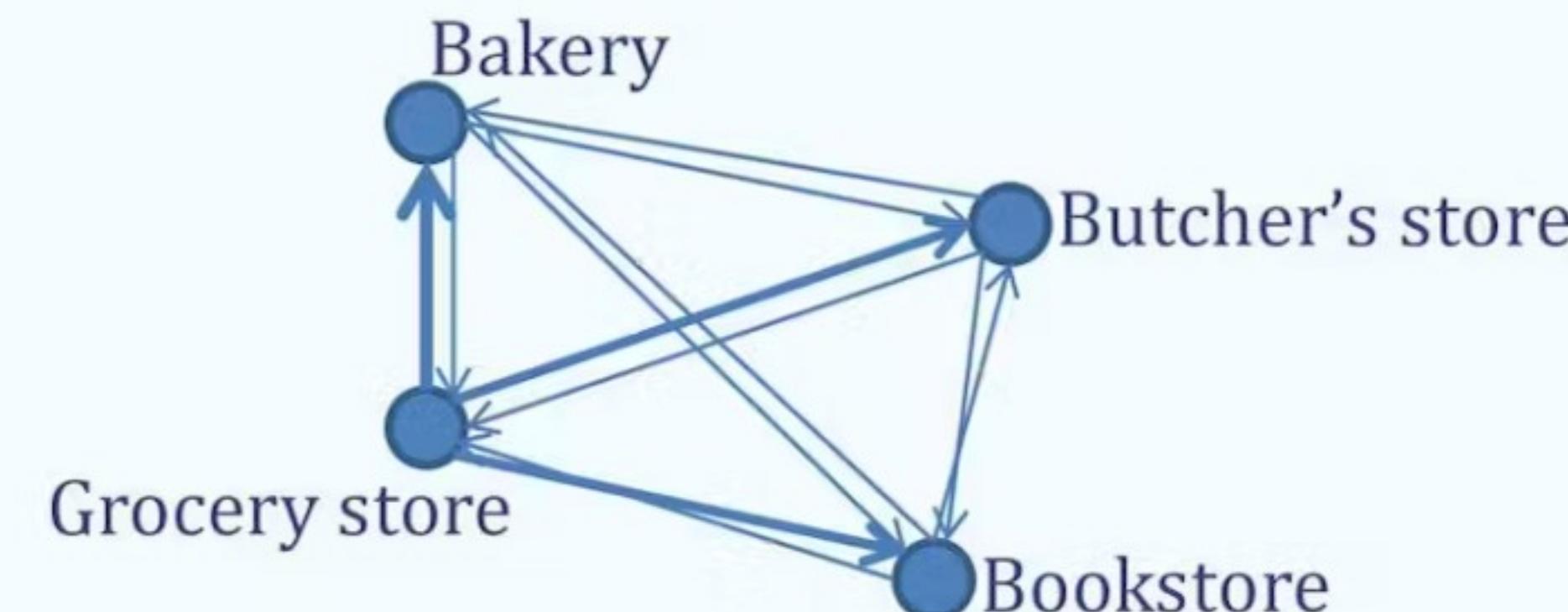


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- Example: Consider an income increase for grocery store owner
- Local autocorrelation:



- Global autocorrelation:



... spatial multiplier process

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- Let's define $z = \lambda W z + \mu$
 - Reduced-form of z is $z = [I - \lambda W]^{-1} \mu$
 - With $\lambda < 1$
- A Leontief expansion yields:
 - $[I - \lambda W]^{-1} = I + \lambda W + \lambda^2 W^2 + \lambda^3 W^3 + \dots$
- $W^2 \rightarrow$ There is an impact of neighbours of neighbours (as defined in W) although it is smaller (λ^2)
 - Global autocorrelation
 - Spatial multiplier process
 - In practice: covariance may approach zero after a relatively small number of powers

What happens when $\lambda > 1$ in $\mathbf{z} = \lambda \mathbf{Wz} + \mu$?

11 ✓

0 ✗

There is no strong positive global autocorrelation

The system will be unstable and will explode

0 ✗

The system will converge very quickly to a new equilibrium

0 ✗

There is a negative global autocorrelation

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- Let's define $z = \lambda W\mu + \mu$
 - This is already a reduced-form of z
- No impact of behaviour beyond 'bands' of neighbours
 - Dependent on definition of W
 - ...Local autocorrelation
- Covariance is zero beyond these bands

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- Local or global autocorrelation?
 - Dependent on application
 - Theory...

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- **Taxonomy:**

$$\mathbf{y} = \rho W\mathbf{y} + \mathbf{X}\boldsymbol{\beta} + \mathbf{W}\mathbf{X}\boldsymbol{\gamma} + \boldsymbol{\epsilon} \quad (1)$$

with

$$\boldsymbol{\epsilon} = \lambda W\boldsymbol{\epsilon} + \boldsymbol{\mu} \quad (2)$$

W is a row-standardised weight matrix
 $\rho, \gamma, \beta, \lambda$ are parameters to be estimated

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- **Taxonomy:**

$$\boxed{y} = \rho W y + \boxed{X\beta} + W X \gamma + \epsilon \quad (1)$$

with

$$\epsilon = \lambda W \epsilon + \mu \quad (2)$$

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- **Taxonomy:**

$$\mathbf{y} = \boxed{\rho W \mathbf{y}} + \mathbf{X}\boldsymbol{\beta} + \mathbf{W}\mathbf{X}\boldsymbol{\gamma} + \boldsymbol{\epsilon} \quad (1)$$

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- **Taxonomy:**

$$\mathbf{y} = \rho W\mathbf{y} + X\boldsymbol{\beta} + \boxed{WX\gamma} + \epsilon \quad (1)$$

with

$$\epsilon = \lambda W\epsilon + \mu \quad (2)$$

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- **Taxonomy:**

$$\mathbf{y} = \rho \mathbf{W} \mathbf{y} + \mathbf{X} \boldsymbol{\beta} + \mathbf{W} \mathbf{X} \boldsymbol{\gamma} + \boldsymbol{\epsilon} \quad (1)$$

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$$\boldsymbol{\epsilon} = \boxed{\lambda \mathbf{W} \boldsymbol{\epsilon}} + \boldsymbol{\mu} \quad (2)$$

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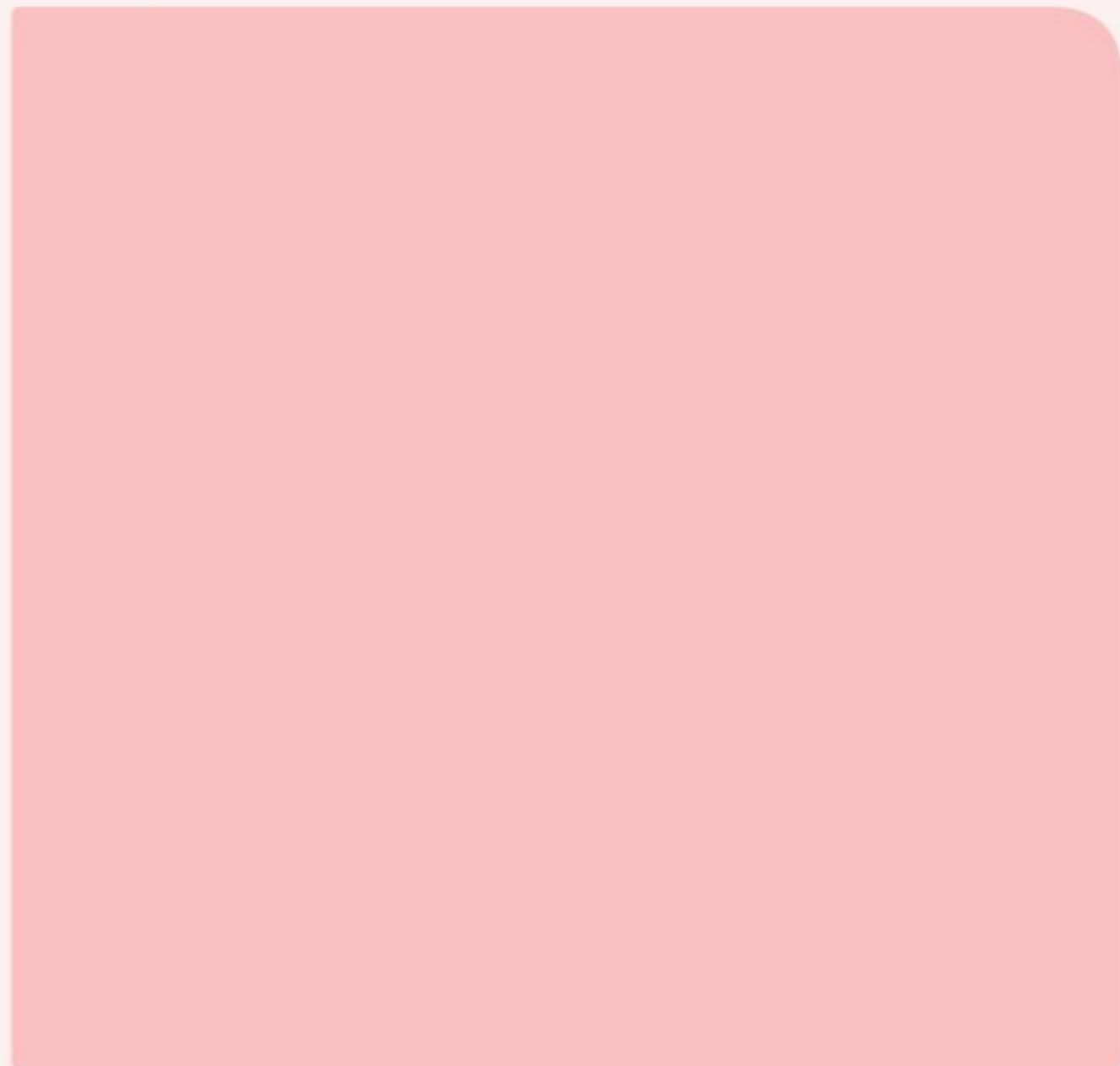
- **Spatial lag model**
 - $y = \rho W y + X\beta + \mu$ (3)
 - $\rho \neq 0, \gamma = 0, \lambda = 0$
 - **Spatial dependence in dependent variables**
- **Note similarity with time-series models**
 - AR Model
 - $y_t = \rho y_{t-1} + X_t \beta + \mu_t$ (4)

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- **Spatial lag model**
 - $y = \rho W y + X\beta + \mu$ (3)
- **The outcome variable influences everyone (indirectly)**
 - Global autocorrelation
- **We may write**
 $(I - \rho W)y = X\beta + \epsilon$
 $y = (I - \rho W)^{-1}(X\beta + \mu)$ with
 $(I - \rho W)^{-1} = I + \rho W + \rho^2 W^2 + \rho^3 W^3 + \dots$

Can the spatial lag model $\mathbf{y} = \rho \mathbf{W}\mathbf{y} + X\beta + \mu$ be estimated by OLS?

9 



0 

Yes, no problem

2 

No, this is not possible

Yes, but the estimator may give the wrong standard errors

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- **Spatial lag model**
 - $y = \rho W y + X\beta + \mu$ (3)
- **We cannot estimate this by OLS because of reverse causality**
- **Recall AR-model:**
$$y_t = \rho y_{t-1} + X\beta + \mu_t$$
 (4)
 - **We can estimate this in principle by OLS because y_{t-1} is not influenced by y_t**

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- **Spatial lag model**
 - **Estimate with OLS?**
 - Let's simplify (3) to
$$\mathbf{y} = \rho W \mathbf{y} + \boldsymbol{\mu} \quad (3')$$
 - **The estimator for ρ yields:**
$$\hat{\rho}_{OLS} = \frac{(\mathbf{W}\mathbf{y})' \mathbf{y}}{(\mathbf{W}\mathbf{y})' (\mathbf{W}\mathbf{y})}$$
- Show that $\hat{\rho}_{OLS}$ is biased when $\text{cov}(\mathbf{y}, \boldsymbol{\mu}) \neq 0$

Consider estimating $\mathbf{y} = \rho \mathbf{W}\mathbf{y} + \mu$ by OLS. Show that ρ_{OLS} is biased when $\text{cov}(\mathbf{y}, \mu) \neq 0$.

0
I am ready!

0
I am stuck...



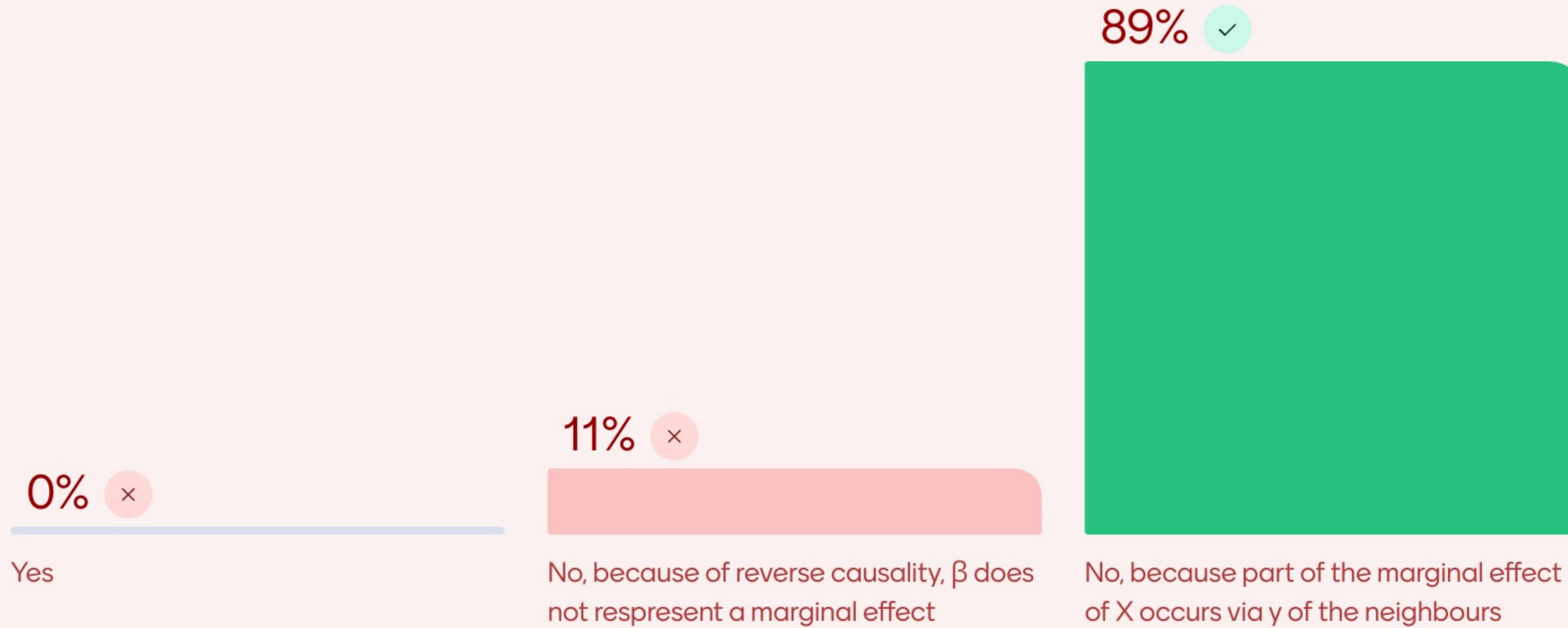
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- **Spatial lag model**
- **Estimate with OLS?**
 - Let's simplify (3) to
$$\mathbf{y} = \rho W\mathbf{y} + \boldsymbol{\mu} \tag{3'}$$
- **The estimator for ρ yields:**
$$\hat{\rho}_{OLS} = \frac{(\mathbf{W}\mathbf{y})'\mathbf{y}}{(\mathbf{W}\mathbf{y})'(\mathbf{W}\mathbf{y})}$$
- **If we plug-in (3') we get:**
$$\hat{\rho}_{OLS} = \frac{(\mathbf{W}\mathbf{y})'(\rho\mathbf{W}\mathbf{y} + \boldsymbol{\mu})}{(\mathbf{W}\mathbf{y})'(\mathbf{W}\mathbf{y})}$$
$$\hat{\rho}_{OLS} = \rho + \frac{(\mathbf{W}\mathbf{y})'\boldsymbol{\mu}}{(\mathbf{W}\mathbf{y})'(\mathbf{W}\mathbf{y})}$$
- **Hence, when $\text{cov}(\mathbf{y}, \boldsymbol{\mu}) \neq 0$, $\hat{\rho}_{OLS}$ is biased**

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- **Spatial lag model**
- **Use maximum likelihood (ML) estimator**
 - **Selects the set of values of the model parameters that maximizes the likelihood function**
- **Instrumental variables (IV)**
 - **Instruments for y may be WX and W^2X^2**
 - **Less efficient than ML, but feasible for 'large' datasets**
 - **e.g. Kelejian and Prucha (1998)**

Assume you use Maximum Likelihood. Does β represent a marginal effect in a spatial lag model $\mathbf{y} = \rho \mathbf{W} \mathbf{y} + X\beta + \mu$?



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- **Spatial cross-regressive model**

- $y = X\beta + \gamma WX + \mu$ (5)
- $\rho = 0, \gamma \neq 0, \lambda = 0$

Can the spatial cross-regressive model $\mathbf{y} = \mathbf{X}\beta + \gamma\mathbf{WX} + \mu$ be estimated by OLS?

6 ✓



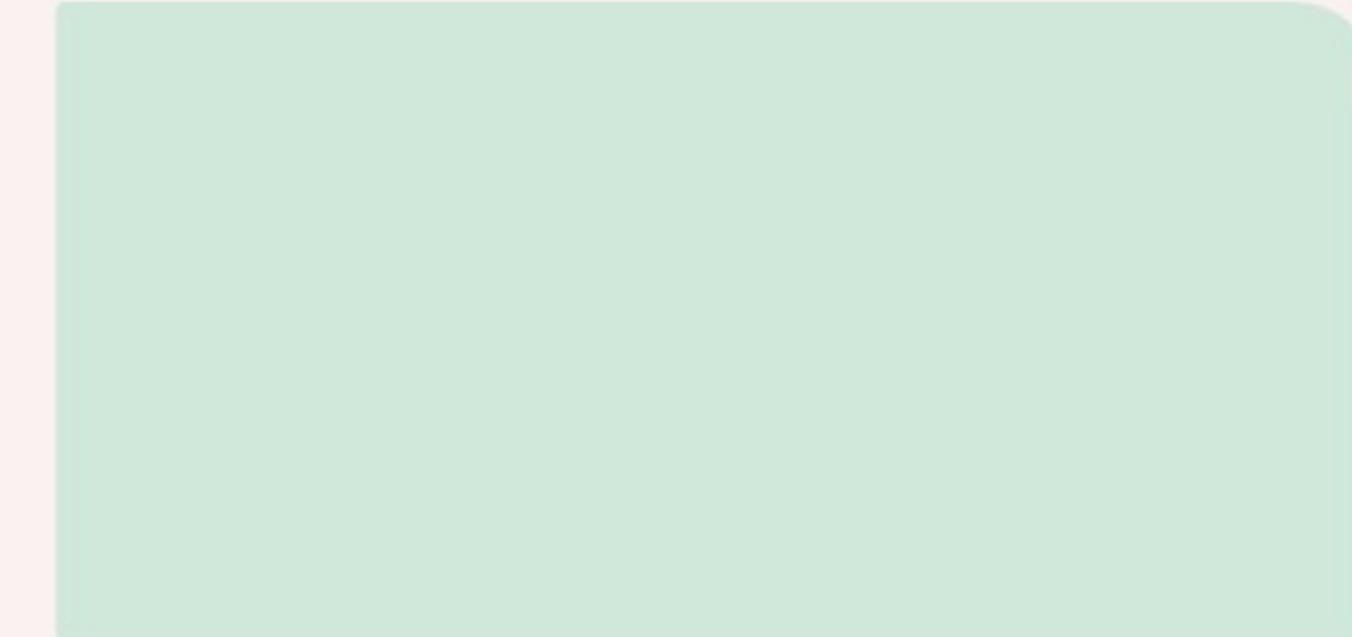
Yes, no problem

3 ✗



Yes, but the estimator may give the wrong standard errors

3 ✗



No, this is not possible

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- **Spatial cross-regressive model**

- $y = X\beta + \gamma WX + \mu$ (5)

- **Include (transformations) of exogenous variables in the regression**
 - OLS is fine!

- **Autocorrelation is local**

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- **Spatial error model**

- $y = X\beta + \epsilon$, with $\epsilon = \lambda W\epsilon + \mu$ (6)
- $\rho = 0, \gamma = 0, \lambda \neq 0$

Can the spatial error model $\mathbf{y} = \mathbf{X}\beta + \lambda\mathbf{W}\epsilon + \mu$ be estimated by OLS?

0

Yes, no problem

10



Yes, but the estimator may give the wrong standard errors

1

No, this is not possible

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- **Spatial error model**
 - $y = X\beta + \epsilon$, with $\epsilon = \lambda W\epsilon + \mu$ (6)
- **Omitted spatially correlated variables**
 - e.g. Ad-hoc defined boundaries
 - Uncorrelated to X !
- **Consistent estimation of parameters β**
- **But: inefficient**
 - ϵ are not i.i.d.
 - Standard errors are higher in OLS
 - β may be different in ‘small’ samples

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- How to apply these models in practice?
 - SPAUTOREG **in STATA**
 - SPATREG **in STATA**
 - **GeoDa (free software, also for large datasets)**
 - PACE'S SPATIAL STATISTICS TOOLBOX **in MATLAB**

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Today:

- **Test spatial autocorrelation using Moran's I**
- **Local vs. global spatial autocorrelation**
- **Incorporate space in regression framework**
- **Spatial regressions**
 - **Spatial lag model**
 - **Spatial cross-regressive model**
 - **Spatial error model**

Spatial econometrics (2)

Applied Econometrics for Spatial Economics

Hans Koster

Professor of Urban Economics and Real Estate